

# Multifractal Spectra of Branching Measure on a Galton-Watson Tree

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## Abstract

If  $Z$  is the branching mechanism for a supercritical Galton-Watson tree with a single progenitor and  $E[Z \log Z] < \infty$ , there is a branching measure  $\mu$  defined on  $\partial\Gamma$  the set of all path  $\xi$  which has a unique node  $\xi|n$  at each generation  $n$ . We use the natural metric  $\rho(\xi, \eta) = e^{-n}$  where  $n = \max\{k : \xi|k = \eta|k\}$  and observe that the local dimension index

$$d(\mu, \xi) = \lim_{n \rightarrow \infty} \frac{\log \mu(B(\xi|n))}{-n} = \alpha = \log m, \quad \mu - a.e. \xi.$$

Our objective is to consider the exceptional points where the above display may fail. There is a non-trivial “thin” spectrum for  $\bar{d}(\mu, \xi)$  when  $p_1 = P\{Z = 1\} > 0$  and  $Z$  has finite moments of all positive orders. Because  $\underline{d}(\mu, \xi) = \alpha$  for all  $\xi$ , we obtain a “thick” spectrum by introducing the “right” power of a logarithm. In both cases we find the Hausdorff dimension of the exceptional sets.

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# 1 Introduction

We are interested in supercritical Galton-Watson trees with a single progenitor. Let  $Z = \{p_0, p_1, \dots\}$  be the offspring distribution of a Galton-Watson branching process, defined on a probability space  $(\Omega, P)$ , with mean  $m = \sum j p_j > 1$  and  $\Gamma = \Gamma(\omega)$  denote the associated family tree. Let  $\mu$  be the branching measure on the boundary  $\partial\Gamma$ ; see Section 2 for more detailed descriptions. We remark that  $\mu$  is a random measure on the (random) tree  $\Gamma$  and our object in this paper is to find properties of the multifractal spectra of  $\mu$  which are true with probability one. Put

$$\alpha = \log m > 0.$$

It is already known that, with probability one,

$$d(\mu, \xi) = \lim_{n \rightarrow \infty} \frac{\log \mu B(\xi|n)}{n} = -\alpha \quad \mu - a.e. \quad \xi \in \partial\Gamma,$$

which can be translated, using the natural metric in  $\partial\Gamma$ , to

$$d(\mu, \xi) = \lim_{r \rightarrow 0} \frac{\log \mu B(\xi, r)}{\log r} = \alpha, \quad \mu - a.e. \quad \xi.$$

The above display is the usual starting point for multifractal analysis of a locally finite Borel measure.

In a recent paper Liu[L3] showed that, if  $p_0 = p_1 = 0$  and  $Z$  has finite moments of all orders then, with probability one,

$$d(\mu, \xi) = \alpha, \quad \forall \xi.$$

Thus the ordinary multifractal spectrum is trivial in this case. However, even in this case, as was shown for the occupation measure of stable subordinator in Shieh-Taylor[ST] and of Brownian motion in Dembo-Peres-Rosen-Zeitouni[DPRZ], one can observe a non-trivial spectrum for “thick” points, by introducing an appropriate power of a logarithm. Results of this type are obtained in Section 4.

In [L3], Liu also observed that, if  $Z$  has finite moments of all positive orders, then

$$\underline{d}(\mu, \xi) = \liminf_{n \rightarrow \infty} \frac{-\log \mu B(\xi|n)}{n} = \alpha \quad \forall \xi \in \partial\Gamma,$$

which implies that for  $\beta \neq \alpha$  the set

$E_\beta = \{\xi \in \partial\Gamma : d(\mu, \xi) = \beta\}$  is empty,

so that the standard multifractal formalism cannot yield a spectrum. However, as in Perkins–Taylor[PT] for super-Brownian motion and Hu–Taylor[HT] for stable occupation measure, we prove that, when  $0 < p_1 < 1$ , there is a non-trivial spectrum for

$$\bar{d}(\mu, \xi) = \limsup_{n \rightarrow \infty} \frac{-\log \mu B(\xi|n)}{n}.$$

In Section 5, we obtain the Hausdorff dimension of

$$C_\beta = \{\xi : \bar{d}(\mu, \xi) \geq \beta\},$$

$$D_\beta = \{\xi : \bar{d}(\mu, \xi) = \beta\},$$

for an interval of values of  $\beta$  in which these sets are non-empty. In both Sections 4 and 5, our method also yields the packing dimension of the relevant sets.

In this paper we again make use of the strong spherical porosity conditions first defined in [PT]. Section 2 defines this condition and its meaning on  $\partial\Gamma$ ; in addition we recall the necessary preliminary definitions and results for Galton–Watson trees. In Dembo–Peres–Rosen–Zeitouni[DPRZ], and Khoshnevisan–Peres–Xiao[KPX], it is pointed out that many exceptional sets examined in random phenomena are of limsup type; they provide a useful theorem giving a lower bound for the Hausdorff dimension of such sets. The exceptional sets which we examine now on  $\partial\Gamma$  are again of limsup type. We therefore develop a version of the main theorem in [DPRZ], proved there for Euclidean cubes, which is valid in the context of a Galton–Watson tree. This is done in Section 3, and is used in both Sections 4 and 5.

## 2 Preliminaries

We begin with notation and results for Galton–Watson trees which we need in this paper; these are adapted from Pemantle–Peres[PP]. Let  $Z = \{p_0, p_1, \dots\}$  be the offspring distribution of a Galton–Watson branching process. We assume that  $m := \sum_j j p_j < \infty$  and that  $p_0 = 0, p_1 < 1$ ; thus  $1 < m < \infty$ , that is we are in the supercritical case (As pointed out in [PP], if  $p_0 > 0$  and  $m > 1$  there is positive probability that  $\Gamma$  is finite. All our results remain true, conditioned on the event that  $\Gamma$  is infinite. However we impose  $p_0 = 0$  to eliminate the complication of conditioning). We also assume that  $Z$  is not

a constant, that is  $p_j < 1$ ,  $\forall j$ . Associated with each realization of the process, we have a (random) family tree, called a *Galton–Watson tree*(GWT), which we denote by  $\Gamma = \Gamma(\omega)$ . Let  $\Gamma_n, n = 0, 1, 2, \dots$  be the  $n$ -th level(generation) so that  $\Gamma = \cup_n \Gamma_n$ . Let  $Z_n$  denote the cardinal number of  $\Gamma_n$  and we assume that  $Z_0 = \{\emptyset\}$ (single progenitor). Assume moreover that  $E[Z \log Z] < \infty$ , then the limit

$$W := \lim_{n \rightarrow \infty} \frac{Z_n}{m^n}$$

exists and is finite and positive a.s., see Artheya–Ney[AN]. For a GWT  $\Gamma$  there is associated a natural *boundary*  $\partial\Gamma$ , which is defined as the set of infinite self-avoiding paths from  $\emptyset$  through the tree; we denote by  $\xi$  a generic point in  $\partial\Gamma$ . For  $\sigma \in \Gamma_n$ ,  $|\sigma| = n$  denotes its length and  $B(\sigma) = B(\sigma, r)$ ,  $r = e^{-n}$ , denotes the “ball”  $\{\xi \in \partial\Gamma : \sigma \text{ is the ancestor of } \xi \text{ in } \Gamma_n\}$ . We also use  $\xi|_n$  to denote the ancestor in  $\Gamma_n$  of an  $\xi \in \partial\Gamma$ . We remark that  $\partial\Gamma$  is a compact metric space under the metric  $d(\xi_1, \xi_2) = e^{-n}$ , where  $n = \max\{k : \xi_1|_k = \xi_2|_k\}$ . In this metric, the subtree consisting of the vertex  $\sigma \in \Gamma_n$  and all its descendants is a ball in  $\partial\Gamma$  of diameter of  $e^{-n}$ , as denoted above. Let  $W(\sigma)$  be the shifted  $W$  at the vertex  $\sigma$ , that is

$$W(\sigma) := \lim_{n \rightarrow \infty} \frac{\text{card}\{\eta \in \Gamma_n : \sigma \text{ is the ancestor of } \eta\}}{m^{n-|\sigma|}}.$$

Since  $\Gamma$  is a countable set,  $W(\sigma)$  exists for all  $\sigma \in \Gamma$  with probability one. By assuming the existence of  $W(\sigma)$ , we define *branching measure* on  $\partial\Gamma$  as the unique (random) Borel measure  $\mu$  on  $\partial\Gamma$  such that

$$\mu B(\sigma) = m^{-n} W_\sigma, \quad \sigma \in \Gamma_n.$$

Note that  $W(\sigma), \sigma \in \Gamma_n$ , are iid with the same distribution as  $W$ , conditional on  $Z_j, j \leq n$ . Thus, the above display reflects the statistical self-similarity of the measure  $\mu$ .

In Perkins–Taylor[PT], the notion of a  $\gamma$ –*thin* set for  $\gamma > 1$  was defined. The definition was for  $\mathbb{R}^d$  but it translates to any metric space, so we define it for  $\partial\Gamma$ .

**Definition 2.1** Fix  $\gamma > 1$ . We say  $E \subset \partial\Gamma$  is  $\gamma$ –*thin* at  $\xi$  if there is a sequence  $r_i \downarrow 0$  such that

$$E \cap [B(\xi, r_i) \setminus B(\xi, r_i^\gamma)] = \emptyset \quad \forall i.$$

We say  $E$  is a  $\gamma$ –*thin* set if  $E$  is  $\gamma$ –*thin* at each  $\xi \in E$ .

The nature of our metric on  $\partial\Gamma$  then asserts that  $\partial\Gamma$  is  $\gamma$ -thin at  $\xi$  if and only if

$$B(\xi|n_i) = B(\xi|[\gamma n_i])$$

for an increasing sequence of positive integers  $n_i$  ( $[\cdot]$ : the greatest integer part). This is equivalent to saying that  $\xi|n \in \Gamma_n$  has exactly one descendant in  $\Gamma_{n+1}$  for  $n_i \leq n \leq [\gamma n_i]$ ,  $i = 1, 2, \dots$ . This shows that  $\partial\Gamma$  can have  $\gamma$ -thin points only if  $p_1 > 0$ , and in this case we denote the set of all  $\gamma$ -thin points for  $\partial\Gamma$  by  $S_\gamma$ .

We now recall some facts first proved in [PT] for Euclidean space. It is easy to check that the results remain true for  $\partial\Gamma$ .

**Lemma 2.1** *For any  $\gamma$ -thin set  $E \subset \partial\Gamma$ , we have*

$$\text{Dim}E \geq \gamma \dim E.$$

Here,  $\text{Dim}$  stands for packing dimension and  $\dim$  stands for Hausdorff dimension. One can refer to [PP] for the detailed definitions and properties of  $\text{Dim}$  and  $\dim$  on GWT's.

**Lemma 2.2** *Let  $\nu$  be any Borel measure on  $\partial\Gamma$  and its support be  $S = S(\nu)$ . If  $A \subset S$  is  $\gamma$ -thin, and*

$$\underline{d}(\nu, x) \geq a \quad \forall x \in A,$$

then,

$$\bar{d}(\nu, x) \geq \gamma a \quad \forall x \in A.$$

We note that, in the case of branching measure  $\mu$ , with probability one, every ball  $B(\sigma)$  has positive  $\mu$  measure. Hence  $S(\mu) = \partial\Gamma$ . We will see that there is a range of values of  $\gamma$  for which  $S_\gamma \neq \emptyset$ , provided some simple conditions are satisfied.

We mention that the metric space  $\partial\Gamma$  has fractal dimension  $\alpha$ : with probability one,

$$\dim \partial\Gamma = \text{Dim} \partial\Gamma = \alpha.$$

### 3 Limsup fractals on Galton–Watson Trees

The following two propositions are a version of Theorem 2.1 of [DPRZ] for GWT. Firstly we note that  $\partial\Gamma$  can be regarded as a random subset of the infinite sequence space  $\mathbb{N}^\infty$ .

We define a random mapping  $\Psi$  on  $\mathbb{N}^\infty \times \Omega$  so that  $\Psi(B, \omega) = 0$ , whenever  $B \not\subset \partial\Gamma(\omega)$ , and that  $Z(B, \omega)$  is  $\{0, 1\}$ -valued, when  $B \subset \partial\Gamma(\omega)$ . Set

$$A = \limsup_n A(n),$$

where

$$A(n) = \cup_{\Psi(B(\sigma))=1, \sigma \in \Gamma_n} B(\sigma).$$

**Proposition 3.1** *Assume that*

(i) *the random variables  $\Psi(B(\sigma))$ ,  $\sigma \in \Gamma_n$ , are independent;*

(ii) *the expectation*

$$q_n := E[\Psi(B(\sigma))] = P\{\Psi(B(\sigma)) = 1\}$$

*is the same for all  $\sigma \in \Gamma_n$ , and*

(iii) *there is a sequence of positive integers  $n_k$  which increase to  $\infty$  rapidly enough so that  $m^{2^{n_k}} \leq n_{k+1}$ ,  $\forall k$ , such that  $q_n$  satisfies the following lower bound estimate,*

$$cm^{-\delta n_k} \leq q_{n_k} \quad \forall k,$$

*where  $c$  is some absolute constant and  $0 < \delta < 1$ . Then, with probability one, the limsup set  $A$  defined above has infinite Hausdorff  $\phi$ -measure,  $\phi-m(A) = +\infty$ , where the gauge function  $\phi$  is defined by*

$$\phi(x) = x^{(1-\delta)\alpha} (\log(1/x))^b, \quad 0 < x < 1, \quad b > 2.$$

*In particular,  $\dim A \geq (1 - \delta)\alpha$ .*

**Proposition 3.2** *Under the conditions of Proposition 3.1, with probability one,*

$$\text{Dim} A = \alpha.$$

Propositions 3.1 and 3.2 can be proved using the methods of [DPRZ]. We remark that we do not need a condition which bounds the correlation since the sub-trees  $B(\sigma)$ ,  $\sigma \in \Gamma_n$ , are completely independent. We need modifications because we are now working on a random tree, rather than the binary tree; these can be made by suitable use of conditional independence. Moreover, the third condition in Proposition 3.1 is stated as required for all large  $n$  in [DPRZ], yet in fact it is only needed for a sufficiently rapidly increasing sequence. We will leave the details to readers. We also mention that Theorem 2.1 of [DPRZ] has been further refined in [KPX].

## 4 Dimension Spectrum for thick behavior

From Section 1, we know that the "typical behavior" of the branching measure  $\mu$  on  $B(\sigma), \sigma \in \Gamma_n$ , is  $m^{-n}$ , for all  $n$  large enough. To describe the behavior of  $\mu$  which is "thicker", we introduce the following two (random) sets

$$A_\theta := \left\{ \xi \in \partial\Gamma : \limsup_n \frac{\mu(B(\xi|n))}{m^{-n}n^\lambda} \geq \theta \right\},$$

and

$$B_\theta := \left\{ \xi \in \partial\Gamma : \limsup_n \frac{\mu(B(\xi|n))}{m^{-n}n^\lambda} = \theta \right\}.$$

In the above,  $\lambda$  is defined by

$$\lambda := 1 - \frac{\alpha}{\log \|Z\|_\infty},$$

where  $\|\cdot\|_\infty$  is the sup norm of the underlying probability space. Note that  $0 < \lambda \leq 1$ , and  $\lambda < 1$  if and only if  $Z$  is a finite distribution. To describe the dimension spectrum of  $A_\theta$  and  $B_\theta$ , we need the following two parameters

$$r_0 := \liminf_{x \rightarrow \infty} \frac{-\log P[W > x]}{x^{1/\lambda}},$$

$$r := \text{the one such that } P[W > x] \approx e^{-rx^{1/\lambda}}, \text{ as } x \uparrow \infty,$$

here  $a \approx b$  means that there are  $c_1, c_2$  such that  $c_1 a \leq b \leq c_2 a$ . To discuss the dimension spectrum of  $A_\theta$ , we assume that  $r_0$  is finite and positive, which is a quite mild assumption as we can see from Lemma 4.1 below. To discuss the dimension spectrum of  $B_\theta$ , we need to impose the stronger assumption that  $r$  exists, and is finite and positive. This is a strong condition, yet it holds for the interesting case that  $Z$  has a geometric distribution, which makes  $W$  have an exponential distribution and then  $\lambda = r = 1$ . It holds also for the case in which  $W$  has a gamma distribution, see Harris[H, p17]. Now we state our dimension spectrum separately for  $A_\theta$  and  $B_\theta$ .

**Theorem 4.1** *Assume that  $r_0$  is finite and positive, then, with probability one,*

$$\dim A_\theta = \alpha - r_0 \theta^{1/\lambda}, \quad 0 \leq \theta \leq \left(\frac{\alpha}{r_0}\right)^\lambda.$$

**Theorem 4.2** *Assume that  $r$  exists, and is finite and positive, then, with probability one,*

$$\dim B_\theta = \alpha - r\theta^{1/\lambda}, \quad 0 \leq \theta \leq \left(\frac{\alpha}{r}\right)^\lambda.$$

Moreover, under the assumption of Theorem 4.1, resp. Theorem 4.2,

$$\begin{aligned} \text{Dim} A_\theta &= \alpha, \quad 0 < \theta < \left(\frac{\alpha}{r_0}\right)^\lambda, \\ \text{resp. } \text{Dim} B_\theta &= \alpha, \quad 0 < \theta < \left(\frac{\alpha}{r}\right)^\lambda. \end{aligned}$$

**Remark 1.** Since  $B_\theta \subset A_\theta$  and  $r$  is necessarily equal to  $r_0$  under the existence assumption, Theorem 4.2 has a stronger assertion under stronger assumption, compared with Theorem 4.1.

**Remark 2.** By the following uniform law for  $\mu$  proved in Liu–Shieh[LS],

$$\limsup_n \sup_{\xi \in \partial\Gamma} \frac{\mu(B(\xi|n))}{m^{-n}n^\lambda} = \left(\frac{\alpha}{r_0}\right)^\lambda,$$

the set  $B_\theta = \emptyset$  for  $\theta > \left(\frac{\alpha}{r_0}\right)^\lambda$ . Thus, the range for  $\theta$  in Theorems 4.1 and 4.2 is exact.

Firstly we state a lemma giving conditions which imply that  $r_0$  is finite and positive. The lemma is a direct consequence of Liu[L1 Theorem 3.1 and L2], however it can be deduced from earlier results.

**Lemma 4.1** *The parameter  $r_0$  is finite and positive, when  $\lambda < 1$ , or when  $\lambda = 1$  and  $E[e^{tZ}] < \infty$  for some, but not for all,  $t > 0$ .*

*Proofs of Theorems.* We concentrate first on the Hausdorff dimension,  $\dim$ . We begin with the discussion of the extreme cases  $\theta = 0$  and  $\theta = \left(\frac{\alpha}{r_0}\right)^\lambda$ . For  $\theta = 0$ , the assertion  $\dim A_\theta = \dim B_\theta = \alpha$  is merely a consequence of the well-known result that  $\dim \mu = \log m$  a.s.; see Hawkes[H] and Lyons–Pemantle–Peres[LPP]. For  $\theta = \left(\frac{\alpha}{r_0}\right)^\lambda$ , it is a consequence of letting a sequence  $\theta_k$  strictly increase to  $\theta$  and proving the spectrum for  $\theta_k$ . Therefore, henceafter we assume that  $\theta$  is not at the endpoints of the range in the theorems.

To prove the upper bound of  $\dim$  it suffices to show that, for any  $b > \alpha - r_0\theta^{1/\lambda}$ , the Hausdorff  $b$ -dimensional measure,  $b - m$ , of  $A_\theta$  is zero. This proof is standard and was given in [LS, Section 3]; we include the proof here for completeness. We observe that, for  $\epsilon : 0 < \epsilon < \theta$  and positive integer  $k$ ,

$$A_\theta \subset \cup_{n \geq k} \left\{ \xi \in \partial\Gamma : \frac{\mu(B(\xi|n))}{m^{-n}n^\lambda} > (\theta - \epsilon)C \right\},$$



where  $C = \left(\frac{\alpha}{r_0}\right)^\lambda$ . We consider the pre-Hausdorff  $b$ -dimensional measure at level  $k$ ,

$$(b-m)_k(A_\theta) = \inf \left\{ \sum_{\sigma} |B(\sigma)|^b : A \subset \cup B(\sigma), |\sigma| \geq k, \forall \sigma \right\}.$$

Recall that  $|B(\sigma)| = e^{-k}$  when  $\sigma \in \Gamma_k$ . Let  $I_k$  denote the random variable defined by

$$I_k = \sum_{|\sigma|=n} |B_\sigma| \mathbb{1} \left\{ \frac{\mu(B(\xi|n))}{m^{-n}n^\lambda} > (\theta - \epsilon)C \right\},$$

then, using the definition of  $r_0$  we see that

$$EI_k \leq \sum_{n \geq k} e^{-(b-\alpha)n} e^{-(r_0-\epsilon)(\theta-\epsilon)^{1/\lambda} C^{1/\lambda} n},$$

when  $k = k(\epsilon)$  is large enough. The series in the above display is convergent, so that  $I_k$  tends to 0 a.s. as  $k \uparrow \infty$ . Since  $\epsilon$  is arbitrary, we conclude that  $b - m(A_\theta) = 0$ .

To obtain the lower bound of  $\dim$ , let  $r_1$  be such that  $r_0 < r_1$  and  $\theta < \left(\frac{\alpha}{r_1}\right)^\lambda$ . We prove that  $\dim A_\theta \geq \alpha - r_1 \theta^{1/\lambda}$  by proving that  $A_\theta$  has infinite Hausdorff  $\phi$ -measure, where

$$\phi(x) = x^{\alpha - r_1 \theta^{1/\lambda}} (\log(1/x))^b, \quad 0 < x < 1, \quad b > 2.$$

We apply Proposition 3.1 in the following way. Define the random mapping  $\Psi(B, \omega)$ ,  $B \subset \mathbb{N}^\infty$  and  $\omega \in \Omega$ , to be 1, only when  $B = B(\sigma)$ ,  $\sigma \in \partial\Gamma_n(\omega)$  and when  $W(\sigma, \omega) > n^\lambda \theta$ ; otherwise  $\Psi$  takes value 0. Thus  $q_n = E[\Psi(B(\sigma))] = P[W > n^\lambda \theta]$ . By our definition of  $r_0$  as a liminf and our choice of  $r_1$  there exists a sequence  $n_k \uparrow \infty$  such that  $q_n \geq e^{-r_1 n \theta^{1/\lambda}}$ ,  $\forall n = n_k$ . We may well assume that  $n_k$  satisfies the rapid growth condition in Proposition 3.1. Therefore Theorem 4.1 is an application of Proposition 3.1 with  $\delta = \alpha - r_1 \theta^{1/\lambda}$  there, and that  $r_1$  can be arbitrarily close to  $r_0$ . To prove the lower bound for  $B_\theta$ , we need to use a strategy first used in [PT]. Under the stronger assumption on the existence of  $r$ , let now

$$\phi(x) = x^{\alpha - r \theta^{1/\lambda}} (\log(1/x))^b, \quad 0 < x < 1, \quad b > 2.$$

Then the above arguments assert that Hausdorff  $\phi$ -measure of  $A_\theta$  is infinite while, by the upper bound proof given in the above, that of  $A_{\theta+1/k}$  is 0 for all  $k = 1, 2, \dots$ . Thus,  $B_\theta = A_\theta \setminus \cup_k A_{\theta+1/k}$  is also of infinite Hausdorff  $\phi$ -measure; in particular we get the desired lower bound for  $B_\theta$ .

Application of Proposition 3.2 gives the assertion for Dim.  $\square$

**Remark.** We believe that the stronger assumption in Theorem 4.2 for  $B_\theta$  is not needed; a proof would require a version of the limsup theorem based on Corollary 3.3 of [KPX], in which the target set satisfies a regularity condition so that we can use the value of  $q_n$  in Proposition 3.1 on a subsequence. We have not formulated a precise theorem, so in this paper we use the methods of [DPRZ].

## 5 Multifractal spectrum for thin behavior

The reader is reminded that smaller than usual branching behavior is reflected in large values of  $\bar{d}(\mu, \xi)$  defined in Section 1. As we have seen in Section 1 that  $d(\mu, \xi)$ , and hence so is  $\bar{d}(\mu, \xi)$ , is equal to  $\alpha$  for all  $\xi$  whenever  $p_0 = p_1 = 0$ ; thus there is only a trivial multifractal spectrum in this case. In this section, we prove that, whenever  $0 < p_1$  there is an interesting spectrum for  $\bar{d}(\mu, \xi)$  (we always assume that  $p_0 = 0$  and that  $p_j < 1, \forall j$ ). We need the following lemma which is in Liu[L3].

**Lemma 5.1** ([L3, Theorem 4.1(ii)]) *If  $Z$  has finite moments of all positive orders, then, with probability one,*

$$\underline{d}(\mu, \xi) = \alpha, \quad \forall \xi \in \partial\Gamma.$$

We introduce a new parameter needed in this section, *assuming that  $p_1 > 0$ .*

$$\tau = -\frac{\log p_1}{\alpha}.$$

Note that

$$p_1 = e^{-\tau\alpha}.$$

The following small tail distribution of  $W$  is known from Bingham[B, p. 217].

$$(5.1) \quad P[W \leq x] \approx x^\tau, \quad x \downarrow 0.$$

We first observe that the probability of a long string of vertices giving rise to a single branch leads to the same estimate for  $P[W \leq x]$ . For, if  $k > n$  and  $\sigma \in \Gamma_n$ , then the event  $E$  that there is only one  $\eta \in \Gamma_k$  descended from  $\sigma$  is

$$(5.2) \quad P(E) = p_1^{k-n}.$$

Now conditional on  $E$ ,  $W_\eta = m^{k-n}W_\sigma$ , so that

$$P[W_\sigma \leq m^{-(k-n)}|E] = P[W_\eta \leq 1].$$

Thus, we have

**Lemma 5.2** *Under the above conditions, if  $x \approx m^{-(k-n)}$ , and  $E$  is the event that each vertex starting with  $\sigma \in \Gamma_n$  has only one descendant up to  $\eta \in \Gamma_k$ , then*

$$1 \geq \frac{P[\{W_\sigma \leq x\} \cap E]}{P(E)} \geq c_3.$$

We now see how to obtain an efficient cover for points in the thin spectrum.

**Lemma 5.3** *Suppose  $\gamma > 1, 0 < \epsilon < (\gamma - 1)/3$ . Then, with probability one, there is an  $n_0 = n_0(\omega)$  such that every vertex  $\sigma \in \Gamma_n$  with  $n \geq n_0$  such that  $W_\sigma \leq e^{-(\gamma-1)\alpha n}$  has fewer than  $e^{\epsilon\alpha n}$  descendants at the level  $k = [(\gamma - \epsilon)n]$ .*

*Proof.* For each  $\eta \in \Gamma_k$  descended from  $\sigma \in \Gamma_n$  such that  $W_\sigma \leq e^{-(\gamma-1)\alpha n}$ , it is seen that

$$W_\eta \leq e^{-\epsilon\alpha n}.$$

By (5.1), the above has probability  $\leq c_2 e^{-\epsilon\alpha\tau n}$ . Putting  $N_\sigma$  as the number of descendants of  $\sigma$  at level  $k$ , we have then

$$P[N_\sigma > e^{\epsilon\alpha n} | W_\sigma \leq e^{-(\gamma-1)\alpha n}] \leq [c_2 e^{-\epsilon\alpha\tau n}]^{e^{\epsilon\alpha n}},$$

in which we have used the fact that  $W_\eta$  for distinct  $\eta \in \Gamma_k$  are iid. Recall that  $Z_n$  counts the vertices  $\sigma \in \Gamma_n$ , thus the expected number of  $\sigma$  such that  $W_\sigma \leq e^{-(\gamma-1)\alpha n}$  and  $N_\sigma > e^{\epsilon\alpha n}$  is

$$\leq E[Z_n] \cdot c_2 e^{-\epsilon\alpha\tau n \cdot e^{\epsilon\alpha n}}.$$

Since  $E[Z_n] = e^{\alpha n}$  we deduce that the probability that there is at least one such vertex is bounded by  $c_2 e^{\alpha n(1 - \epsilon\tau \cdot e^{\epsilon\alpha n})}$ , which is the general term of a convergent series. By the Borel–Cantelli Lemma we have proved the lemma.  $\square$

**Lemma 5.4** *If  $0 < p_1 < 1$ , then, with probability one,*

$$(5.3) \quad \dim C_\beta \leq \alpha \left[ \frac{\alpha}{\beta} (1 + \tau) - \tau \right], \quad \alpha \leq \beta,$$

where  $C_\beta = \{\xi \in \partial\Gamma : \bar{d}(\mu, \xi) \geq \beta\}$ .

When  $\beta > \alpha(1 + 1/\tau)$ , the right hand side of (5.3) is negative, and we interpret this as stating that  $C_\beta = \emptyset$ .

*Proof.* (i) When  $\beta = \alpha$ , (5.3) is immediate.

(ii) When  $\beta > \alpha(1 + 1/\tau)$ , we will prove that  $C_\beta = \emptyset$  a.s. Take  $\zeta$  such that  $\beta > \zeta > \alpha(1 + 1/\tau)$ , then  $\bar{d}(\mu, \xi) \geq \beta$  implies that  $\mu B(\xi|n) < e^{-\zeta n}$  for infinitely many integers  $n$ . The expected number of those  $\sigma \in \Gamma_n$  for which  $W_\sigma < e^{(\alpha-\zeta)n}$  is

$$E[Z_n] \cdot P[W < e^{(\alpha-\zeta)n}] = e^{(\alpha+(\alpha-\zeta)\tau)n},$$

which is a negative power of  $e^n$ . By Borel-Cantelli Lemma, we deduce that, for  $n \geq n_1 = n_1(\omega)$

$$\mu B(\xi|n) \geq e^{-\zeta n}, \quad \forall \xi.$$

By letting  $\beta \downarrow \alpha(1 + 1/\tau)$  through a countable set, we deduce that, with probability one,

$$\bar{d}(\mu, \xi) \leq \alpha(1 + 1/\tau), \quad \forall \xi \in \partial\Gamma.$$

(iii) Now suppose that  $\alpha < \beta < \alpha(1 + 1/\tau)$ . Put  $\gamma = \frac{\beta}{\alpha} > 1$ . Instead of covering the vertices  $\sigma \in \Gamma_n$  where  $\mu B(\xi|n) < e^{-\beta n}$  by balls of diameter  $e^{-n}$  we use the descendant vertices at level  $k = [(\gamma - \epsilon)n]$  which can be covered by balls of diameters  $e^{-k}$ . By Lemma 5.3, when  $n$  is large enough the number of such vertices is less than  $e^{\epsilon \alpha n}$  so that the total number needed will be at most  $e^{\epsilon \alpha n} \cdot T_n$  where  $T_n$  is the number of those  $\sigma \in \Gamma_n$  for which  $W_\sigma < e^{-(\beta-\alpha)n}$ . Now  $E[T_n] = m^n \cdot e^{-(\beta-\alpha)\tau n}$ , so that we obtain

$$E[s^\delta - m(C_\beta)] \leq \sum_{n=n_1}^{\infty} e^{-[(\gamma-\epsilon)n]\delta + \epsilon \alpha n - (\beta-\alpha)\tau n + \alpha n},$$

where  $n_1$  is arbitrary. If the power of  $e^n$  in this series is negative, we deduce that  $s^\delta - m(C_\beta) = 0$  a.s. This will be true if

$$\delta > \delta_\epsilon := \frac{1 + \epsilon - (\gamma - 1)\tau}{\gamma - \epsilon} \cdot \alpha.$$

Letting  $\epsilon \downarrow 0$  through a countable set, we see that  $s^\delta - m(C_\beta) = 0$  a.s. for  $\delta \geq \frac{(1-(\gamma-1)\tau)\alpha}{\gamma}$ . Substituting  $\gamma = \beta/\alpha$  we prove the assertion.  $\square$ .

We are now ready to prove that (5.3) gives the right answer for  $\dim C_\beta$ . However, if we are to obtain the same answer for  $\dim D_\beta$ , as in Section 4, we need to find a gauge function  $\phi(s) = s^\Delta L(s)$  with  $L(s)$  slowly varying, such that  $\phi - m(C_\beta) = \infty$ . We will prove this by applying Proposition 3.1, and the strategy is the same as that used in [PT]: we find a random Cantor-like subset  $T_\gamma$  which is  $\gamma$ -thin and use Proposition 3.1 to find

its Hausdorff  $\phi$ -measure. This set  $T_\gamma \subset C_\beta$  by Lemma 2.2 and Lemma 5.1. In order to apply Proposition 3.1, fix a  $\gamma > 1$ , we define the random mapping  $\Psi$  there by defining  $\Psi(B, \omega) = 1$  if and only if  $B = B(\sigma)$ ,  $\sigma \in \Gamma_n(\omega)$  is such that its ancestor in  $\Gamma_{[n/\gamma]-1}(\omega)$  has a string of single branches stretching to  $\sigma$ . Denote the limsup set  $A$  there now by  $T_\gamma$ . By (5.2), the probability  $q_n$  in Proposition 3.1 is now

$$q_n \geq c \cdot p_1^{(1-1/\gamma)n} \cdot e^{\alpha(1-1/\gamma)n} = c \cdot e^{\alpha n(1-1/\gamma)(\tau+1)}.$$

We remark that the third factor in the middle term of the above display comes from the expected number of all the possible ancestors in  $\Gamma_{[n/\gamma]-1}$ , given an element in  $\Gamma_n$ . We can now calculate the  $\delta$  in Proposition 3.1. Note that  $T_\gamma$  is clearly a  $\gamma$ -thin subset of  $\partial\Gamma$  by the construction. Thus we have

**Lemma 5.5** *Assume that  $0 < p_1 < 1$  and  $Z$  has finite moments of all positive orders; let  $\beta$  be fixed,  $\alpha < \beta < (1 + 1/\tau)\alpha$ , and define  $\gamma = \beta/\alpha$ . Then the Hausdorff measure of the  $\gamma$ -thin set  $T_\gamma$  defined above satisfies  $\phi - m(T_\gamma) = +\infty$ , where the gauge function  $\phi$  is  $\phi(x) = x^\Delta (\log(1/x))^\beta$ , with*

$$\Delta = \alpha \left[ \frac{1}{\gamma} (1 + \tau) - \tau \right]$$

We can now state our main decomposition.

**Theorem 5.1** *If  $0 < p_1 < 1$  and  $Z$  has finite moments of all positive orders, then the branching measure  $\mu$  has the following properties, with probability one. Set*

$$C_\beta = \{\xi : \bar{d}(\mu, \xi) \geq \beta\}, D_\beta = \{\xi : \bar{d}(\mu, \xi) = \beta\},$$

then

- (a)  $C_\beta$  and therefore  $D_\beta$  is empty for  $\beta > \alpha(1 + \frac{1}{\tau})$ .
- (b)  $D_\beta$  is non-empty for  $\alpha \leq \beta \leq \alpha(1 + \frac{1}{\tau})$ , and in this range

$$\begin{aligned} \dim C_\beta = \dim D_\beta &= \alpha \left[ \frac{\alpha}{\beta} (1 + \tau) - \tau \right] \\ \text{Dim} C_\beta = \text{Dim} D_\beta &= \alpha. \end{aligned}$$

*Proof.* By Lemmas 2.2 and 5.1,  $T_\gamma \subset C_\beta$ , where  $\gamma = \beta/\alpha$ . By Lemma 5.5, the Hausdorff  $\phi_\Delta$ -measure of  $C_\beta$  is  $+\infty$ , where  $\phi_\Delta$  is the gauge function there in Lemma 5.5. Regard  $\Delta$  as a function of  $\beta$ , it is strictly monotone. Lemma 5.4 then tells that  $\phi_\Delta$  measure of  $C_{\beta+1/k}$  is 0. Thus, arguing as in the proofs of Theorems 4.1 and 4.2, we see that  $\dim D_\beta \geq \Delta$ . Since  $D_\beta \subset C_\beta$ , we have completed the proof for  $\alpha \leq \beta < \alpha(1 + 1/\tau)$ . In the case where  $\beta = \alpha(1 + 1/\tau)$  we only need to show that  $D_\beta = C_\beta$  is non-empty. This will follow if we can construct  $T_\gamma$  for  $\gamma = 1 + 1/\tau$  by requiring the string of single branches to stretch from the level  $[n/\gamma - \log n] - 1$  to  $n$ . This condition forces  $\bar{d}(\mu, \xi) \geq \alpha(1 + 1/\tau)$ , on using Lemmas 2.2 and 5.1.  $\square$

**Remark.** In Theorem 5.1 we assume that  $Z$  has finite moments of all positive orders is mainly to apply Lemma 5.1. It seems possible that we may weaken the condition in Theorem 5.1 to, say, that  $Z$  has finite moments up to a certain order  $k_0$  greater than one and get a spectrum involving  $p_+ = \sup\{a \geq 1 : EZ^a < \infty\}$  (one critical value in [L3]).

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