

# Ultra High Frequency Volatility Estimation with Market Microstructure Noise

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- $\varepsilon$  summarizes a diverse array of **market microstructure effects**, either **informational or not**: bid-ask bounces, discreteness of price changes, differences in trade sizes or informational content of price changes, gradual response of prices to a block trade, the strategic component of the order flow, inventory control effects, etc.

- We study the implications of such a data generating process for the estimation of the volatility of the efficient log-price process

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- In theory, **sampling as often as possible** will produce in the limit a perfect estimate of that quantity.

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- $T^{1/2} (\hat{\sigma}^2 - \sigma^2) \xrightarrow{T \rightarrow \infty} N(0, 2\sigma^4 \Delta)$
- Thus selecting  **$\Delta$  as small as possible** is optimal for the purpose of estimating  $\sigma^2$ .



- When volatility is stochastic,  $dX_t = \sigma_t dW_t$  :
  - **Realized volatility**  $\sum_{i=1}^n (X_{t_{i+1}} - X_{t_i})^2$  estimates the **quadratic variation**  $\int_0^T \sigma_t^2 dt$ .



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  - As in the constant  $\sigma$  case, selecting  $\Delta$  as small as possible ( $= n$  as large as possible) is optimal.

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- Asymptotics are in  $\Delta \rightarrow 0$ , with  $T$  fixed.
- The usual estimator of  $\langle X, X \rangle_T$  is the **realized volatility**

$$[Y, Y]_T = \sum_{i=1}^n (Y_{t_{i+1}} - Y_{t_i})^2.$$

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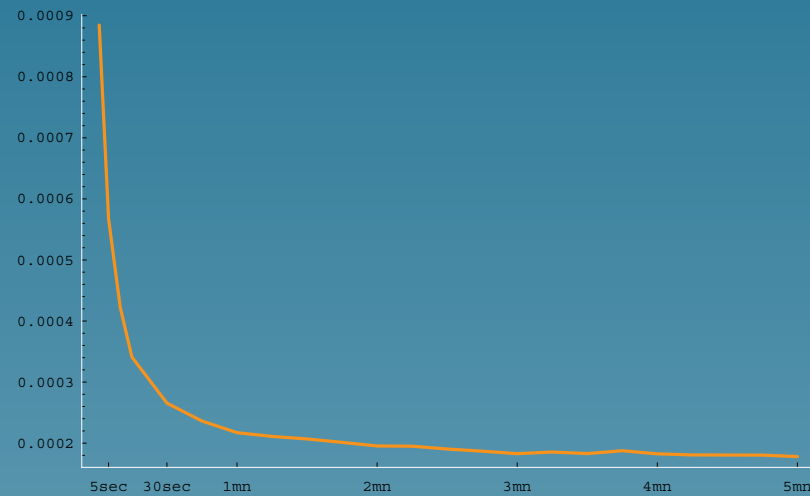
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- Here is the fourth best estimator for different values of  $\Delta$ , averaged for the 30 DJIA stocks and the last 10 trading days in April 2004:



- As  $\Delta = T/n \rightarrow 0$ , the graph shows that the estimator diverges as predicted by our result ( $2nE[\varepsilon^2]$ ) **instead of converging to the object of interest**  $\langle X, X \rangle_T$  as predicted by standard asymptotic theory.

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- This gives rise to the **third best** estimator we define as  $[Y, Y]_T^{(sparse, opt)}$ .

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- Hence the bias of  $[Y, Y]_T^{(avg)}$  can be consistently estimated by  $\frac{\bar{n}}{n}[Y, Y]_T^{(all)}$ .

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- We call this estimator **Two Scales Realized Volatility**.

- We show that if the number of subsamples is optimally selected as  $K^* = cn^{2/3}$ , then **TSRV** has the following distribution:

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 \widehat{\langle X, X \rangle}_T &\stackrel{\mathcal{L}}{\approx} \underbrace{\langle X, X \rangle}_T \\
 &\quad \text{object of interest} \\
 &+ \frac{1}{n^{1/6}} \left[ \underbrace{\frac{8}{c^2} E[\varepsilon^2]^2}_{\text{due to noise}} + \underbrace{c \frac{4T}{3} \int_0^T \sigma_t^4 dt}_{\text{due to discretization}} \right]^{1/2} Z_{\text{total}} \\
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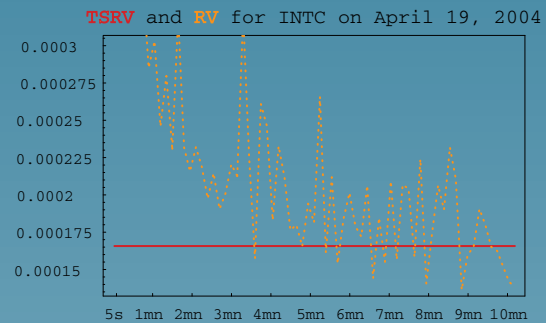
- Unlike all the previously considered ones, this estimator is now **correctly centered**
- To the best of our knowledge, this is the **only consistent estimator** for  $\langle X, X \rangle_T$  in the presence of market microstructure noise.

## 5. Monte Carlo Simulations

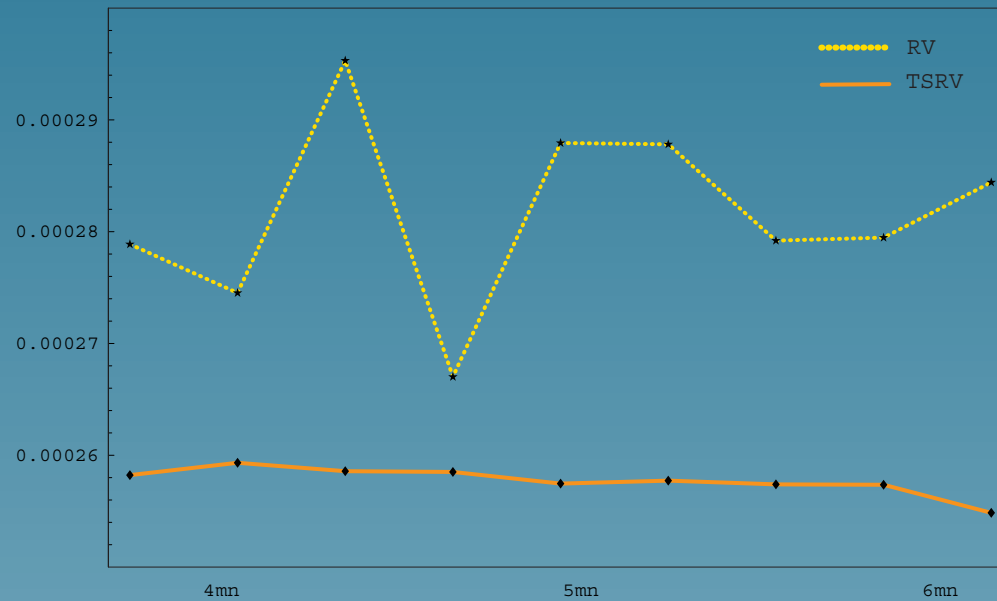
	Fifth Best $[Y, Y]_T^{(all)}$	<b>RV</b> Fourth Best $[Y, Y]_T^{(sparse)}$	Third Best $[Y, Y]_T^{(sparse,opt)}$	Second Best $[Y, Y]_T^{(avg)}$	<b>TSRV</b> First Best $\widehat{\langle X, X \rangle}_T^{(adj)}$
Small Sample Bias Asymptotic Bias	$1.1699 \cdot 10^{-2}$ $1.1700 \cdot 10^{-2}$	$3.89 \cdot 10^{-5}$ $3.90 \cdot 10^{-5}$	$2.18 \cdot 10^{-5}$ $2.20 \cdot 10^{-5}$	$1.926 \cdot 10^{-5}$ $1.927 \cdot 10^{-5}$	$2 \cdot 10^{-8}$ 0
Small Sample Variance Asymptotic Variance	$1.791 \cdot 10^{-8}$ $1.788 \cdot 10^{-8}$	$1.4414 \cdot 10^{-9}$ $1.4409 \cdot 10^{-9}$	$1.59 \cdot 10^{-9}$ $1.58 \cdot 10^{-9}$	$9.41 \cdot 10^{-10}$ $9.37 \cdot 10^{-10}$	$9 \cdot 10^{-11}$ $8 \cdot 10^{-11}$
Small Sample RMSE Asymptotic RMSE	$1.1699 \cdot 10^{-2}$ $1.1700 \cdot 10^{-2}$	$5.437 \cdot 10^{-5}$ $5.442 \cdot 10^{-5}$	$4.543 \cdot 10^{-5}$ $4.546 \cdot 10^{-5}$	$3.622 \cdot 10^{-5}$ $3.618 \cdot 10^{-5}$	$9.4 \cdot 10^{-6}$ $8.9 \cdot 10^{-6}$
Small Sample Relative Bias	182	0.61	0.18	0.15	-0.00045
Small Sample Relative Variance	82502	1.15	0.11	0.053	0.0043
Small Sample Relative RMSE	340	1.24	0.37	0.28	0.065

## 6. Data Analysis

- Here is a comparison of **RV** to **TSRV** for INTC, last 10 trading days in April 2004:



- Zooming around the 5 minutes sampling frequency:





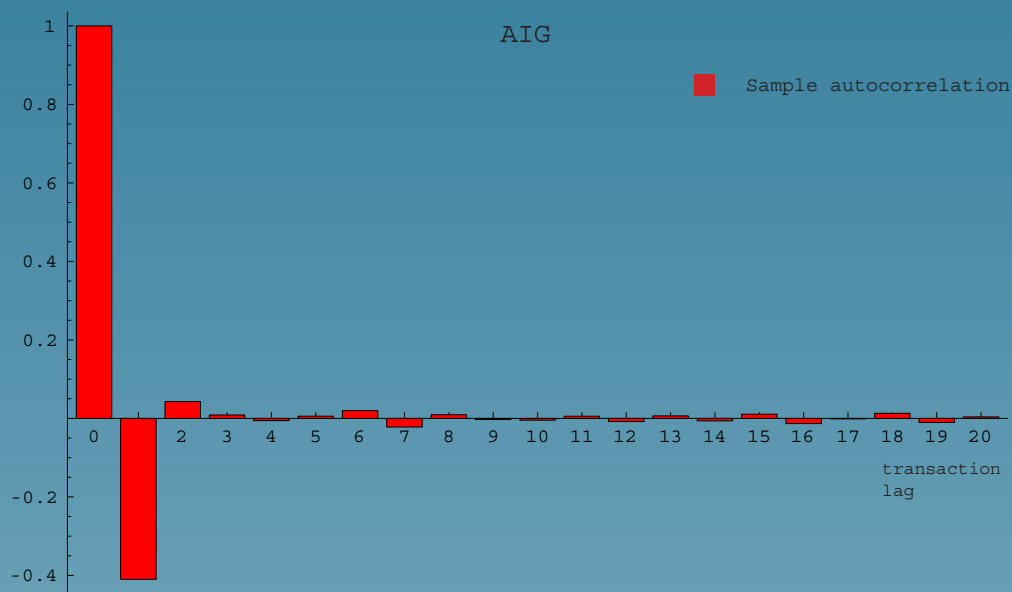
## 7. Dependent Market Microstructure Noise

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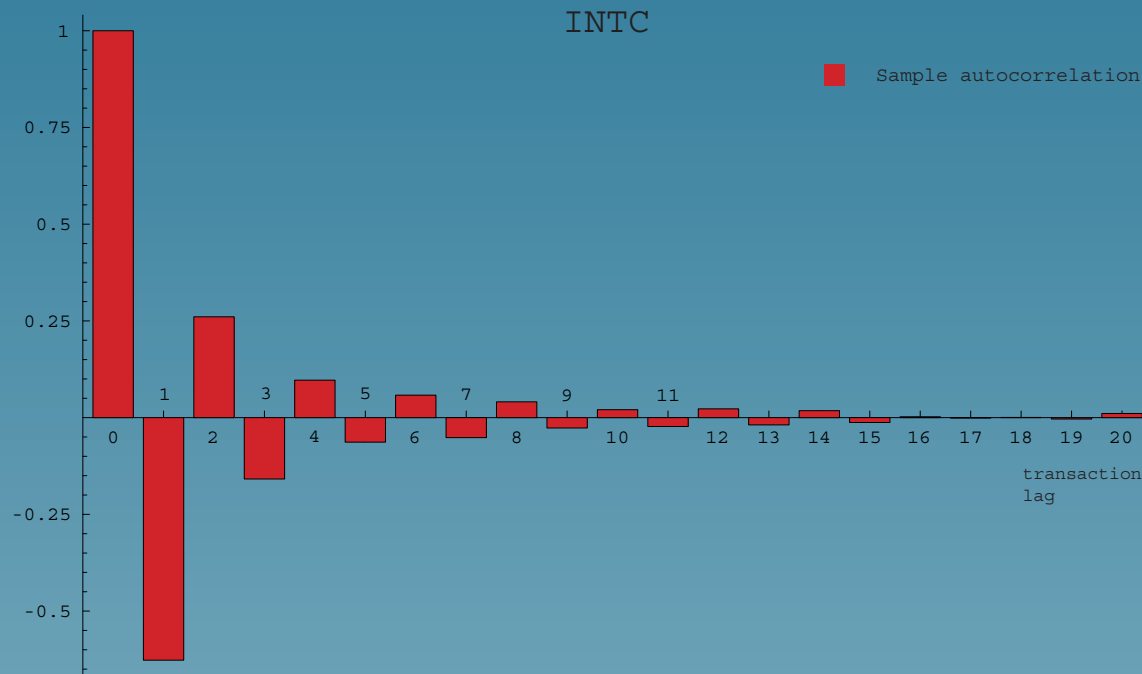
- So far, we have assumed that the noise  $\varepsilon$  was iid.
- In that case, log-returns are **MA(1)**:

$$Y_{\tau_i} - Y_{\tau_{i-1}} = \int_{\tau_{i-1}}^{\tau_i} \sigma_t dW_t + \varepsilon_{\tau_i} - \varepsilon_{\tau_{i-1}}$$

- For example, here is the **autocorrelogram** for AIG transactions, last 10 trading days in April 2004:



- But here is the **autocorrelogram** for INTC transactions, same last 10 trading days in April 2004:



- A simple model to capture this higher order dependence is

$$\varepsilon_{t_i} = U_{t_i} + V_{t_i}$$

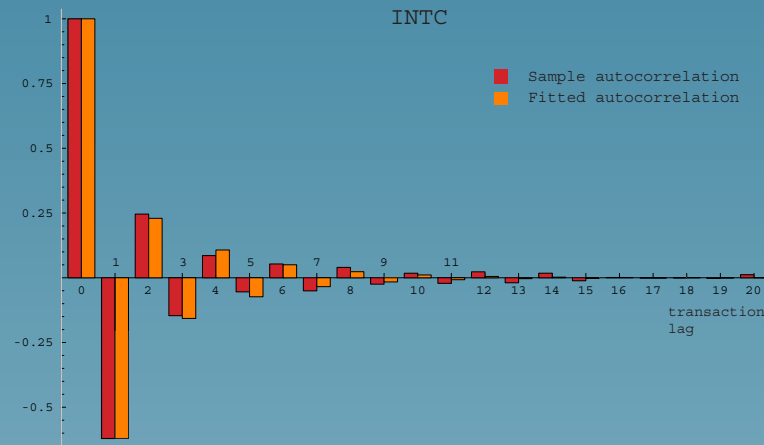
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- Fitted autocorrelogram for INTC:



- The **TSRV** Estimator with  $(J, K)$  Time Scales

$$\widehat{\langle X, X \rangle}_T = \underbrace{[Y, Y]_T^{(K)}}_{\text{slow time scale}} - \frac{\bar{n}_K}{\bar{n}_J} \underbrace{[Y, Y]_T^{(J)}}_{\text{fast time scale}}$$

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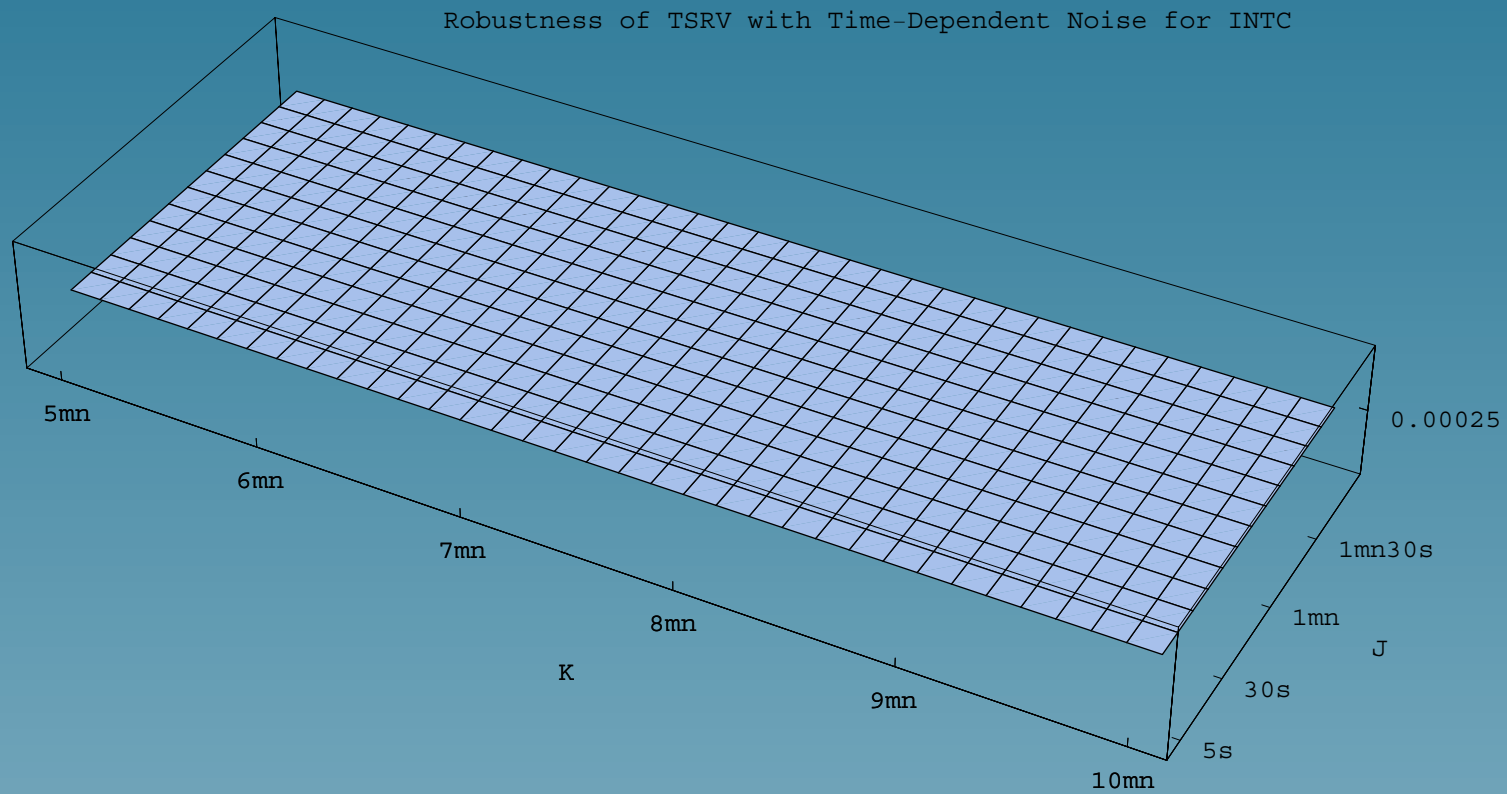


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- Specifically, we let the noise process  $\varepsilon_{t_i}$  be stationary and **strong mixing** with exponential decay. We also suppose that  $E \left[ \varepsilon^{4+\kappa} \right] < \infty$  for some  $\kappa > 0$ .

- Robustness to the selection of the slow ( $K$ ) and fast ( $J$ ) time scales, INTC again:



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- TSRV corresponds to the special case where  $M = 1$ , i.e., where one uses a **single slow time scale** in conjunction with the fast time scale to bias-correct it.

- For suitably selected weights  $a_i$  and  $M = O(n^{1/2})$ ,  $\widehat{\langle X, X \rangle}_T^{(\text{msrv})}$  converges to the  $\langle X, X \rangle_T$  at rate  $n^{-1/4}$ .

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  - No matter what the model is, no matter what quantity is being estimated.