

Mixed models: design of experiments

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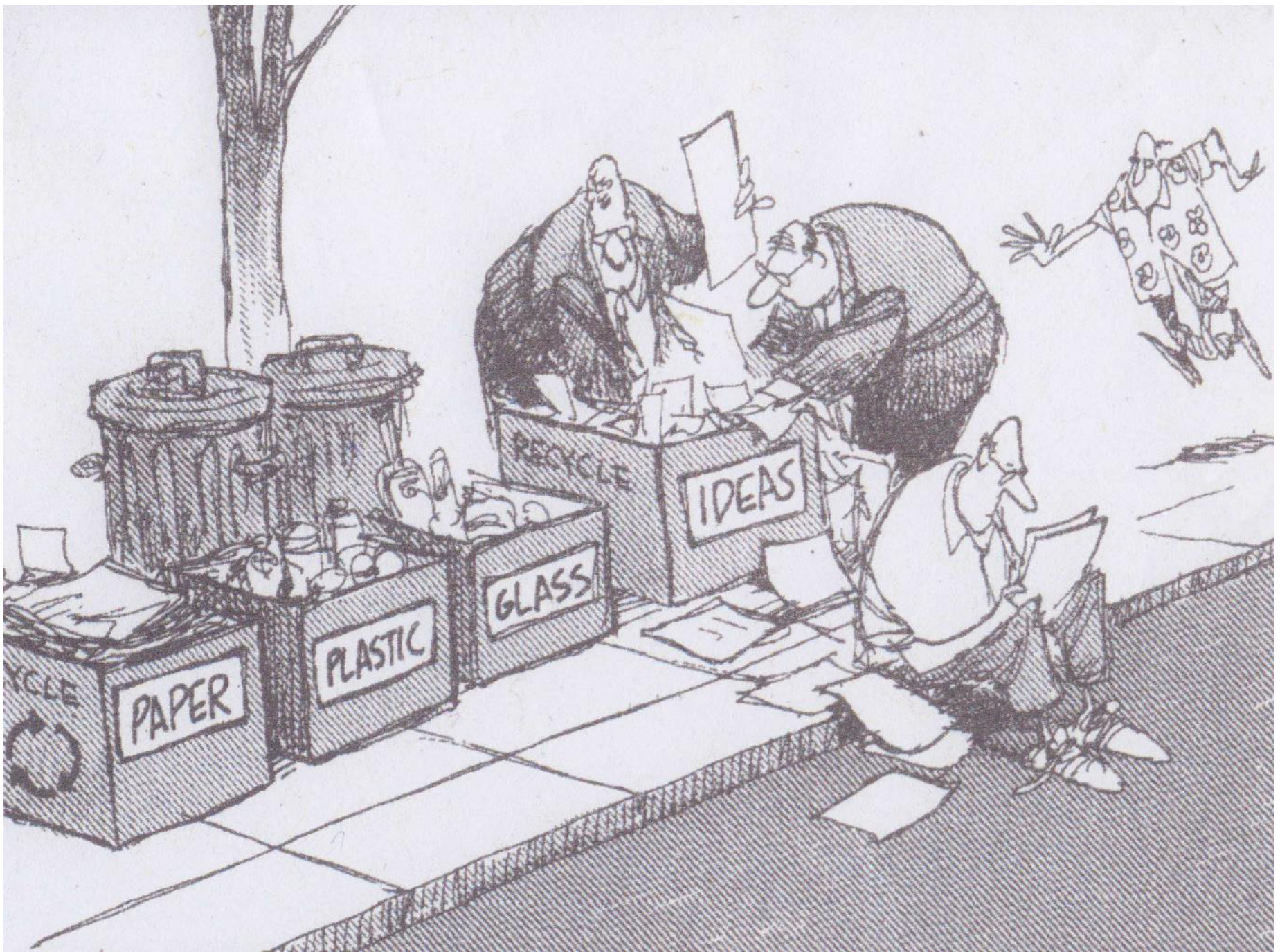
August, 2011

Selected references (estimation)

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Selected references (design)

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Type I mixing

Observations are independent and

Model 1

$$\mathbf{y}_i \sim p(\mathbf{y} | \mathbf{x}_i, \theta)$$

Model 2

$$\mathbf{y}_i \sim p(\mathbf{y} | \mathbf{x}_i, \theta_i)$$

$$\theta_i \sim p(\theta | \Theta)$$

For transition from (2) to (1) use: $\mathbf{y} \sim p(\mathbf{y} | \mathbf{x}, \Theta) = \int p(\mathbf{y} | \mathbf{x}, \theta) p(\theta | \Theta) d\theta$

Model 3

$$E(\mathbf{y}_i) = \eta(\mathbf{x}_i, \theta)$$

$$\text{Var}(\mathbf{y}_i) = \mathbf{S}(\mathbf{x}_i, \theta)$$

Model 4

$$E(\mathbf{y}_i) = \eta(\mathbf{x}_i, \theta_i)$$

$$\text{Var}(\mathbf{y}_i) = \mathbf{V}(\mathbf{x}_i, \theta_i)$$

$$E(\theta_i) = \theta$$

$$\text{Var}(\theta_i) = \Sigma(\theta)$$

For transition from (4) to (3) use:

$$E(\mathbf{y}) = E_{\theta}[E_{\mathbf{y}}(\mathbf{y} | \theta)]$$

$$\text{Var}(\mathbf{y}) = E_{\theta}[\text{Var}_{\mathbf{y}}(\mathbf{y} | \theta)] + \text{Var}_{\theta}[E_{\mathbf{y}}(\mathbf{y} | \theta)]$$

Type II mixing

$$y_{ij} \sim p(\mathbf{y} | \theta_i)$$
$$\theta_i \sim p(\theta | \vartheta(\mathbf{x}_i, \Theta))$$

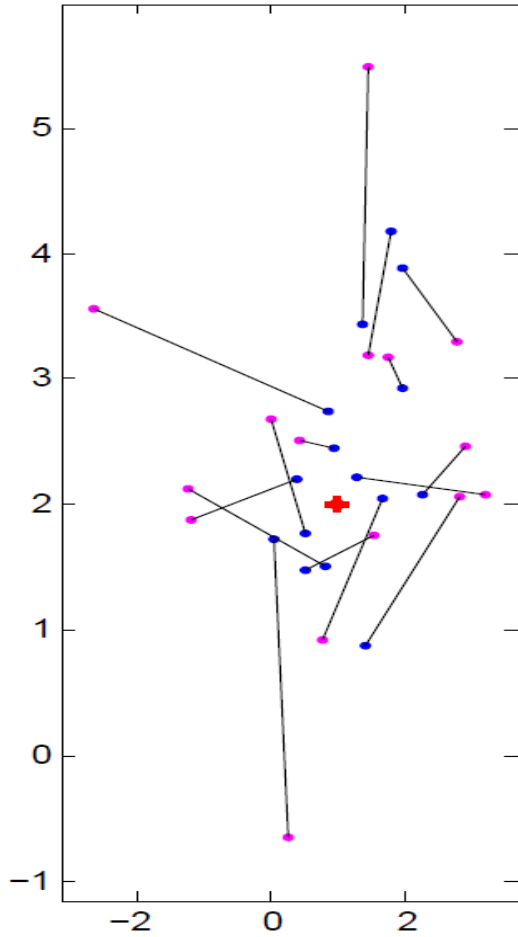
Two major problems

Problem 1: Not much is known about population. A substantial number of observations is needed to understand its distribution.

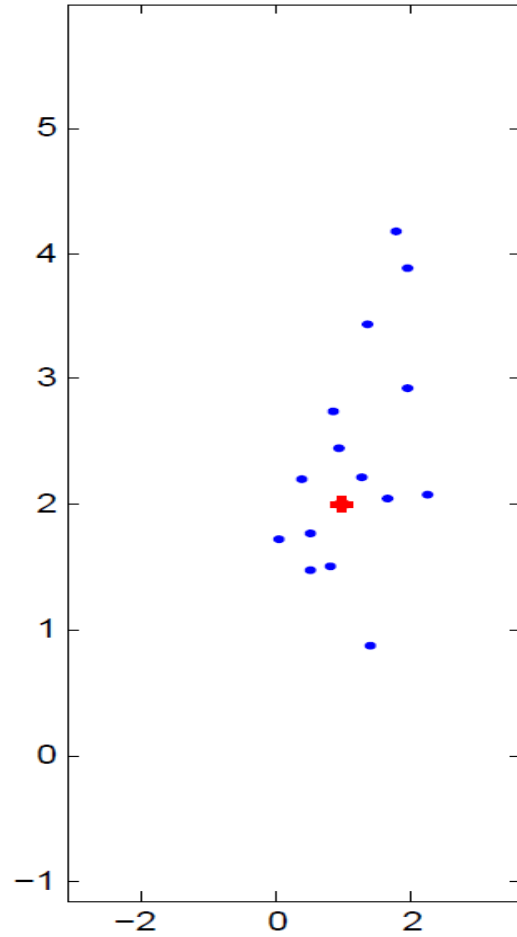
Problem 2: Population is well known, individual parameters for item “i” should be predicted. How to make it with relatively few observations

Learning about population

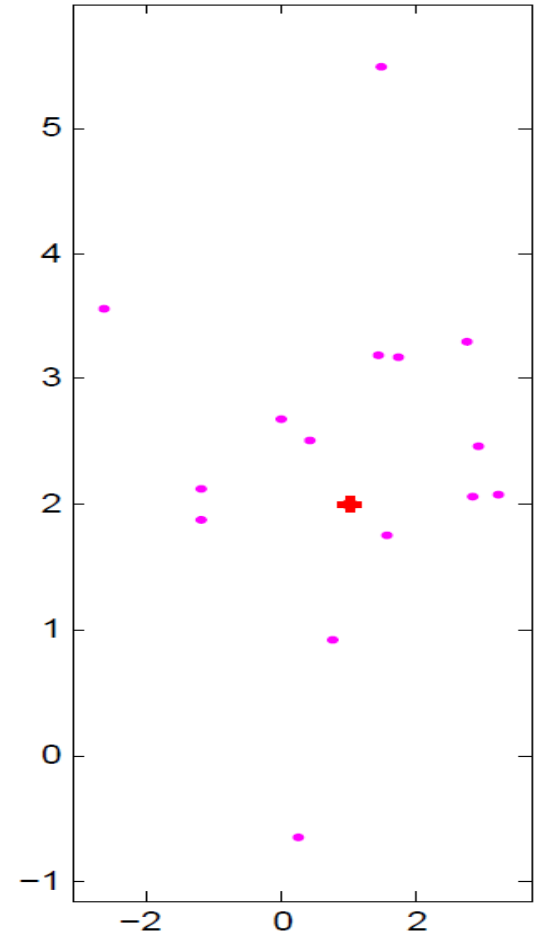
Random effects and observed values



True random effects



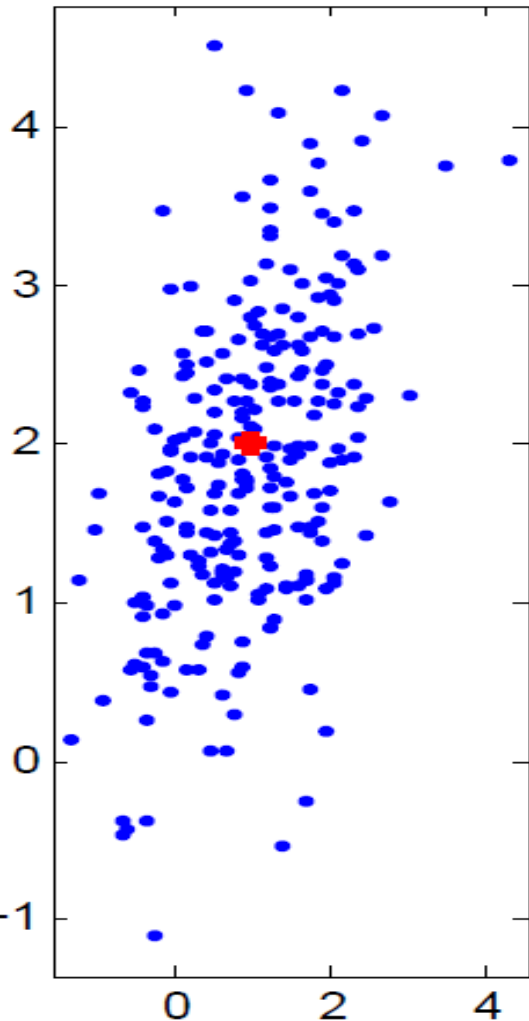
Observed values



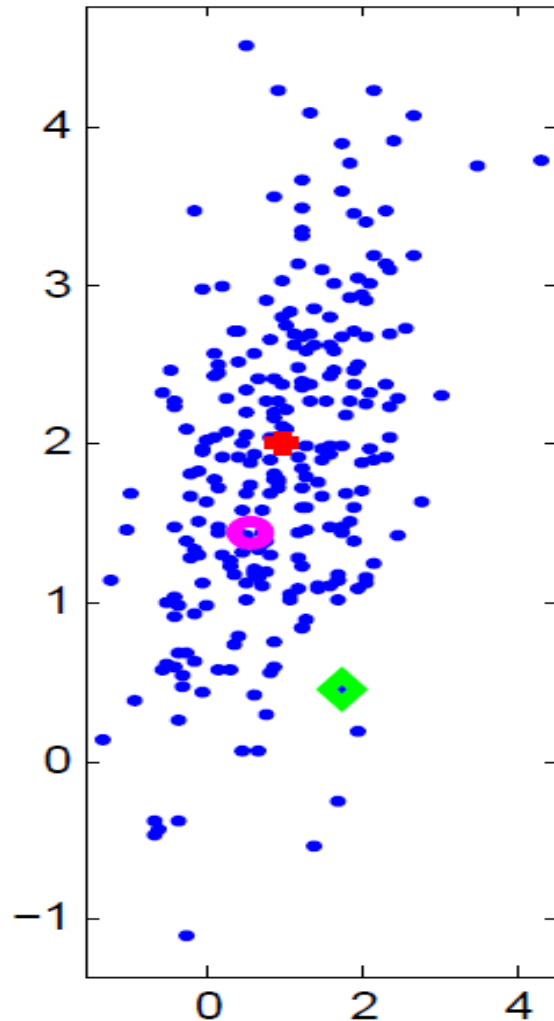
Should we spend more time making more “points” or making shorter “segments”?

Predicting for patients

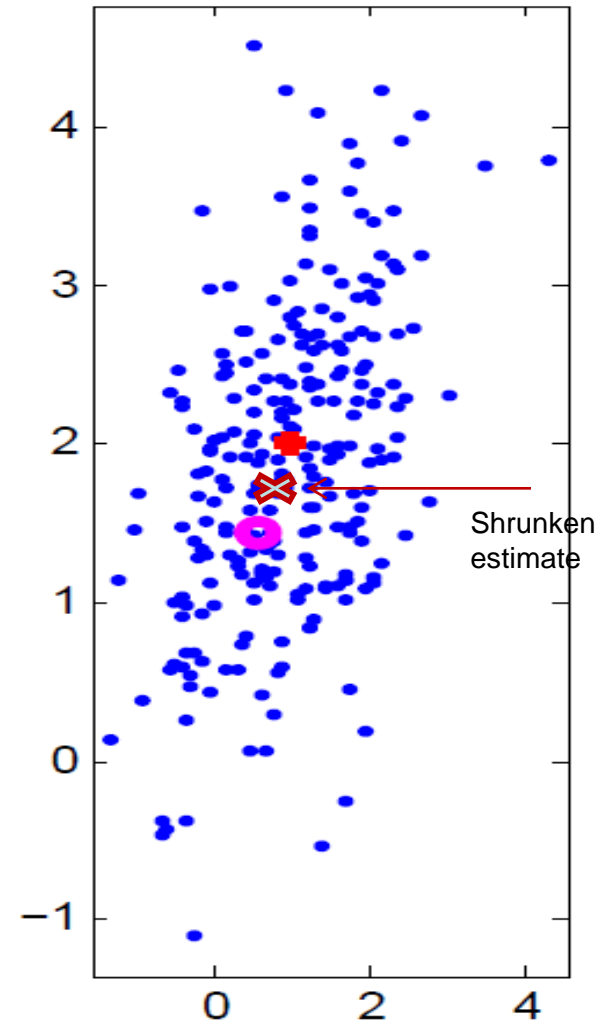
Random effects



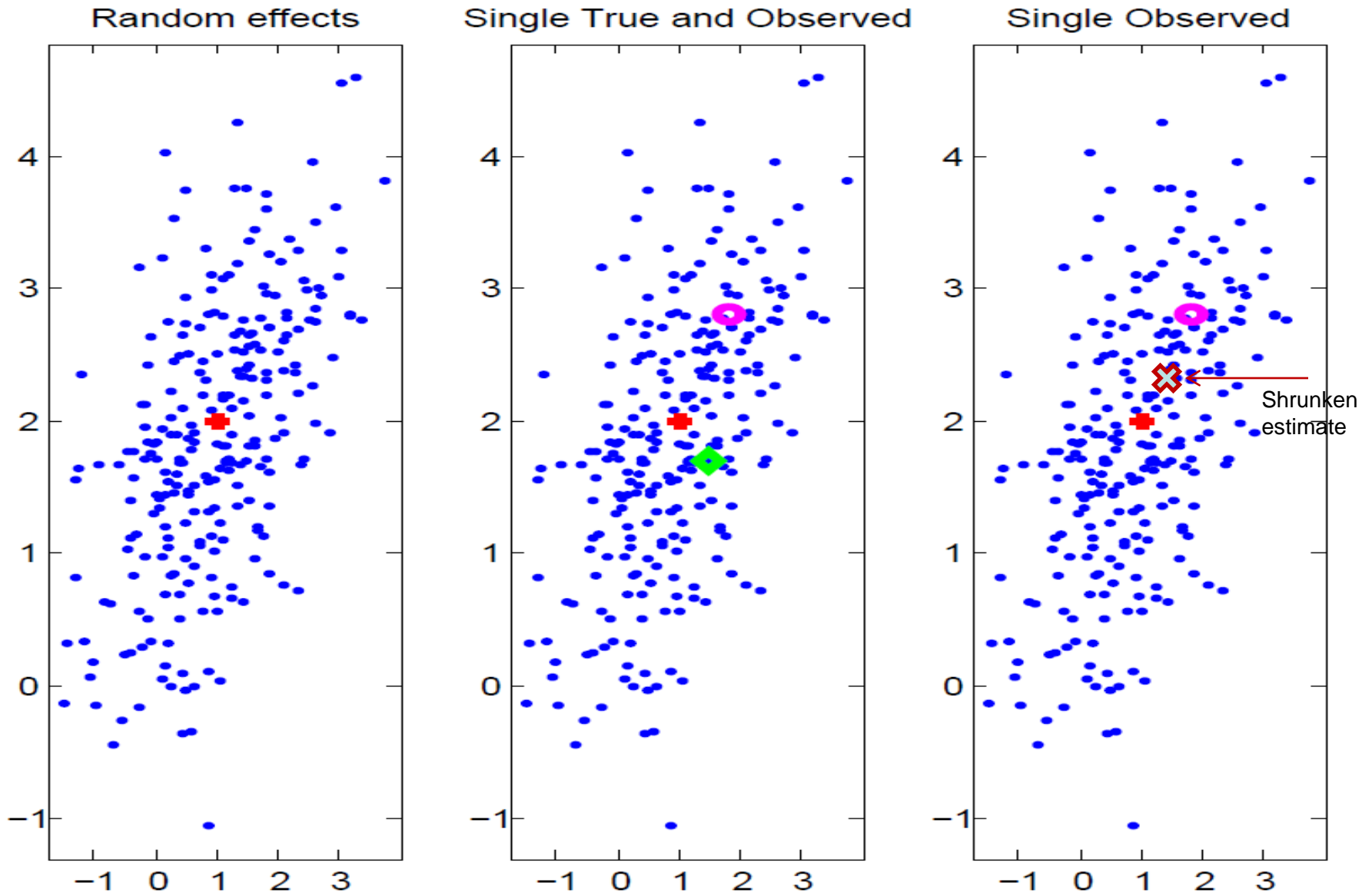
Single True and Observed



Single Observed



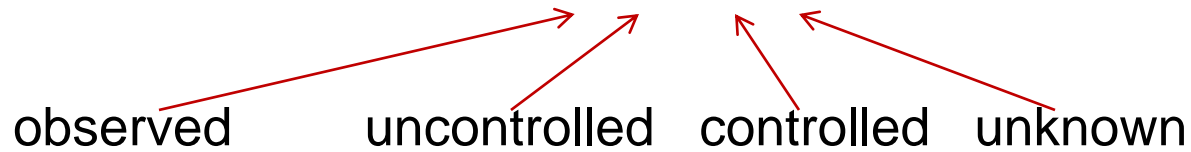
Predicting for patients



Short intro to optimal design

- Model

$$\mathbf{Y} \sim p(\mathbf{y} | \mathbf{t}, \mathbf{x}, \theta)$$



- Utility

$$\mathbf{U}(\mathbf{t}, \mathbf{x}, \theta)$$

- Cost

$$\varphi(\mathbf{t}, \mathbf{x}, \theta)$$

- Design region

$$\mathbf{x} \in \mathcal{X}$$

Estimation: MLE

Likelihood function:

$$\mathcal{L}_N(\theta) = \prod_{i=1}^n \prod_{j=1}^{r_i} p(\mathbf{y}_{ij} | \mathbf{x}_i, \theta)$$

Maximum likelihood estimator:

$$\hat{\theta}_N = \arg \max_{\theta \in \Omega} \prod_{i=1}^n \prod_{j=1}^{r_i} p(\mathbf{y}_{ij} | \mathbf{x}_i, \theta)$$

Asymptotic normality:

$$(\hat{\theta}_N - \theta_0) / \sqrt{N} \sim \mathcal{N}(0, D(\xi, \theta_0))$$

Normalized var-cov matrix:

$$D(\xi, \theta) = M^{-1}(\xi, \theta)$$

Design:

$$\xi = \{\mathbf{x}_i, w_i\}_1^n, \quad w_i = r_i / N$$

Total cost:

$$Cost = N\Phi(\xi) = N \sum_{i=1}^n w_i \varphi(\mathbf{x}_i)$$

Estimation: Information matrix

$$NM(\xi, \theta_0) = \sum_{i=1}^n n_i \mu(\mathbf{x}_i, \theta_0) = N \sum_{i=1}^n w_i \mu(\mathbf{x}_i, \theta_0)$$

Information matrix of a “single” observation:

$$\mu(\mathbf{x}, \theta_0) = E\{\zeta(\mathbf{y}|\mathbf{x}, \theta)\zeta^T(\mathbf{y}|\mathbf{x}, \theta)\}_{\theta=\theta_0}$$

Score function: $\zeta(\mathbf{y}|\mathbf{x}, \theta) = \frac{\partial}{\partial \theta} \log p(\mathbf{y}|\mathbf{x}, \theta)$

For transformed parameters:

$$M(\xi, \theta) = JM(\xi, \vartheta(\theta))J^T$$

$$J = \frac{\partial \vartheta^T}{\partial \theta} = \left\| \frac{\partial \vartheta_\alpha}{\partial \theta_\beta} \right\|_{1,1}^{m',m}$$

Main optimization problem

Optimal design:

$$\begin{aligned} \xi^* &= \arg \min_{\xi \in \Xi(\mathcal{X})} \Psi(N(\xi)M(\xi, \theta)) \\ \text{s.t. } & N(\xi)\Phi(\xi) \leq C \end{aligned}$$

Equivalently:

$$\xi^* = \arg \min_{\xi \in \Xi(\mathcal{X})} \Psi \left(\frac{M(\xi, \theta)}{\Phi(\xi)} \right)$$

The same but with “prior” information

Optimal design:

$$\begin{aligned} \xi^* &= \arg \min_{\xi \in \Xi(\mathcal{X})} \Psi(\mathbf{M}_0 + N(\xi)M(\xi, \theta)) \\ \text{s.t. } & N(\xi)\Phi(\xi) \leq C \end{aligned}$$

Equivalently:

$$\xi^* = \arg \min_{\xi \in \Xi(\mathcal{X})} \Psi \left(\frac{\mathbf{M}_0}{C} + \frac{M(\xi, \theta)}{\Phi(\xi)} \right)$$

Popular optimality criteria

- D-criterion: $\Psi(D) = |D|$
- Linear criterion: $\Psi(D) = \text{tr}AD = \text{tr}LDL^T$
- E-criterion: $\Psi(D) = \lambda_{\max}(D)$

For any design:

$$|D|^{1/m} \leq m^{-1} \text{tr}D \leq \lambda_{\max}(D)$$

$$|D|^{1/m} \leq m^{-1} \text{tr}AD, \quad |A| = 1$$

$$L^T DL \leq \lambda_{\max}(D), \quad L^T L = 1$$

Basic “design” formulae

Necessary and sufficient conditions (D-criterion):

$$\text{tr}[\mu(\mathbf{x}, \theta)M^{-1}(\xi^*, \theta)] \leq m\varphi(\mathbf{x})/\Phi(\xi^*)$$

First order algorithm (D-criterion)

Step forward:

$$\mathbf{x}_{s+1}^{\oplus} = \arg \max_{\mathbf{x} \in \mathcal{X}} \{ \text{tr}[\mu(\mathbf{x}, \theta)M^{-1}(\xi_s, \theta)] - m\varphi(\mathbf{x})/\Phi(\xi_s) \}$$

Step backward:

$$\mathbf{x}_{s+1}^{\ominus} = \arg \min_{\mathbf{x} \in \mathcal{X}_s} \{ \text{tr}[\mu(\mathbf{x}, \theta)M^{-1}(\xi_s, \theta)] - m\varphi(\mathbf{x})/\Phi(\xi_s) \}$$

Note. If there is a prior information then use:

$$\text{tr}\{[\mu(\mathbf{x}, \theta) + \mathbf{M}_0\varphi(\mathbf{x})/C][M(\xi^*, \theta) + \mathbf{M}_0\Phi(\xi^*)/C]^{-1}\} \leq m\varphi(\mathbf{x})/\Phi(\xi^*)$$

Specific cases from 1988, A&F

Table 1: Dual problems for the most used criteria.

$\Psi(M)$	$\varphi(x, \xi)$
$\ln M^{-1} $	$m - \text{tr } M^{-1}m(x)$
$\ln A^T M^{-1}A $	$s - \text{tr } M^{-1}A[A^T M^{-1}A]^{-1}A^T M^{-1}m(x), \quad s = \text{rank } A$
$\text{tr } A^T M^{-1}A$	$\text{tr } A^T M^{-1}A - \text{tr } M^{-1}A A^T M^{-1}m(x)$
$\int_{X_0} w(x) d(x, \xi) dx$ $d(x, \xi) = \text{tr } M^{-1}m(x)$	$\int_{X_0} w(x) d(x, \xi) dx - \int_{X_0} w(x') \text{tr } M^{-1}m(x') M^{-1}m(x) dx'$
$m^{-1}(\text{tr } M^{-p})^{1/p}$	$\text{tr } M^{-p} - \text{tr } M^{-p-1}m(x)$

Optimal Design Construction

Information matrix of a single observation, cost function, design region (candidate points), optimality criterion



Optimal design, comparison of different designs, sample size determination, ...

Linear case with known \mathbf{V} and Σ

$$\begin{aligned} E(\mathbf{y}_i) &= \eta(\mathbf{x}_i, \theta_i) = \mathbf{F}^T(\mathbf{x}_i)\theta_i, & \text{Var}(\mathbf{y}_i) &= \mathbf{V} \\ E(\theta_i) &= \theta, & \text{Var}(\theta_i) &= \Sigma \end{aligned}$$

$$\hat{\theta} = \left[\sum_i [\mathbf{M}^{-1}(\mathbf{x}_i) + \Sigma]^{-1} \right]^{-1} \sum_i [\mathbf{M}^{-1}(\mathbf{x}_i) + \Sigma]^{-1} \bar{\theta}_i$$

$$\mathbf{M}(\mathbf{x}_i) = \mathbf{F}(\mathbf{x})\mathbf{V}^{-1}\mathbf{F}^T(\mathbf{x})$$

$$\bar{\theta}_i = \mathbf{M}^{-1}(\mathbf{x}_i)\mathbf{F}(\mathbf{x}_i)\mathbf{y}_i$$

$$\hat{\theta}_i = [\mathbf{M}(\mathbf{x}_i) + \Sigma^{-1}]^{-1} [\mathbf{M}(\mathbf{x}_i)\bar{\theta}_i + \Sigma^{-1}\theta]$$

$$\hat{\hat{\theta}}_i = [\mathbf{M}(\mathbf{x}_i) + \Sigma^{-1}]^{-1} [\mathbf{M}(\mathbf{x}_i)\bar{\theta}_i + \Sigma^{-1}\hat{\theta}]$$

Estimation of population parameters: information matrix for a “single” observation

$$\mu_{\alpha\beta}(\mathbf{x}, \boldsymbol{\theta}) = \frac{\partial \boldsymbol{\eta}^T(\mathbf{x}, \boldsymbol{\theta})}{\partial \theta_\alpha} \mathbf{S}^{-1}(\mathbf{x}, \boldsymbol{\theta}) \frac{\partial \boldsymbol{\eta}(\mathbf{x}, \boldsymbol{\theta})}{\partial \theta_\beta} + \frac{1}{2} \operatorname{tr} \left[\mathbf{S}^{-1}(\mathbf{x}, \boldsymbol{\theta}) \frac{\partial \mathbf{S}(\mathbf{x}, \boldsymbol{\theta})}{\partial \theta_\alpha} \mathbf{S}^{-1}(\mathbf{x}, \boldsymbol{\theta}) \frac{\partial \mathbf{S}(\mathbf{x}, \boldsymbol{\theta})}{\partial \theta_\beta} \right]$$

Linear case: $\boldsymbol{\eta}(\mathbf{x}, \boldsymbol{\theta}) = \mathbf{F}^T(\mathbf{x}) \boldsymbol{\theta}$

$$\mathbf{S}(\mathbf{x}, \boldsymbol{\Sigma}, \mathbf{V}(\mathbf{x})) = \mathbf{F}^T(\mathbf{x}) \boldsymbol{\Sigma} \mathbf{F}(\mathbf{x}) + \mathbf{V}(\mathbf{x})$$

$$\boldsymbol{\mu}(\mathbf{x}) = \mathbf{F}(\mathbf{x}) [\mathbf{F}^T(\mathbf{x}) \boldsymbol{\Sigma} \mathbf{F}(\mathbf{x}) + \mathbf{V}(\mathbf{x})]^{-1} \mathbf{F}^T(\mathbf{x})$$

If \mathbf{M} is regular then:

$$\boldsymbol{\mu}(\mathbf{x}) = [\mathbf{M}^{-1}(\mathbf{x}) + \boldsymbol{\Sigma}]^{-1}$$

Note: Go to slides 16-17 to learn how to build an optimal design

Estimation of individual parameters

- *The best unbiased prediction for i -th item:*

maximize $|\mathbf{M}(\mathbf{x})|$

$$\begin{aligned}\mu(x) &= \mathbf{f}(x)\mathbf{f}^T(x) \\ \psi(x) &= \text{tr} \mu(x)\mathbf{M}(\xi) - m\end{aligned}$$

- *For the best individual prediction on average:*

maximize $|\mathbf{M}^{-1}(\mathbf{x}) + \Sigma|^{-1}$

$$\begin{aligned}\mu(x) &= \mathbf{f}(x)\mathbf{f}^T(x) \\ \psi(x) &= \text{tr}\{[\mu(\mathbf{x}) + n\Sigma^{-1}][M(\xi^*) + n\Sigma^{-1}]^{-1}\} - m\end{aligned}$$

Unknown variances

Parameters $\Theta = \{\theta, \Sigma, \mathbf{V}\}$, where $\mathbf{V} = \sigma^2 \mathbf{I}$ are unknown.

Information matrix of a “single” observation:

$$\mu(\mathbf{x}, \Theta) = \begin{pmatrix} \mu_{\theta\theta}(\mathbf{x}) & \mathbf{0}_{m,m} & \mathbf{0}_{m,1} \\ \mathbf{0}_{m,m} & \mu_{\Sigma\Sigma}(\mathbf{x}) & \mu_{\Sigma\sigma}(\mathbf{x}) \\ \mathbf{0}_{1,m} & \mu_{\Sigma\sigma}^T(\mathbf{x}) & \mu_{\sigma\sigma}(\mathbf{x}) \end{pmatrix}$$

$$\mu_{\theta\theta}(\mathbf{x}) = \mathbf{F}(\mathbf{x})\mathbf{S}^{-1}(\mathbf{x})\mathbf{F}^T(\mathbf{x})$$

$$\{\mu_{\Sigma\Sigma}(\mathbf{x})\}_{\alpha\beta} = \frac{1}{2} \left[\mathbf{F}_\alpha(\mathbf{x})\mathbf{S}^{-1}\mathbf{F}_\beta^T(\mathbf{x}) \right]^2 \quad \{\mu_{\Sigma\sigma}(\mathbf{x})\}_\alpha = \frac{1}{2} \mathbf{F}_\alpha(\mathbf{x})\mathbf{S}^{-2}\mathbf{F}_\alpha^T(\mathbf{x})$$

$$\mu_{\sigma\sigma}(\mathbf{x}) = \frac{1}{2} \text{tr} \left[\mathbf{S}^{-2} \right]$$

$$\mathbf{F}(\mathbf{x}) = [f(x_1), \dots, f(x_k)]$$

Summary

- What will the mixed effects model be used for?
- Definition of a “single” observation.
- Derivation of its information matrix.
- Understanding the cost structure.
- Computing
- Benchmarking