



Optimal design of blocked and split-plot experiments for fixed-effects and variance-components estimation

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Optimal design of blocked and split-plot **response surface** experiments for fixed-effects and variance-components estimation

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Outline

- blocked and split-plot experiments
- model
- optimal design approach
 - fixed effects
 - variance components
- computational aspects
 - Gauss-Stieltjes-Wigert quadrature
 - Gauss-Jacobi quadrature
- computational results



Blocked experiments

- useful whenever experimental runs cannot be performed under homogeneous circumstances
- runs are partitioned in groups or blocks
- runs within a block are more alike than runs from different blocks
- blocks can be days, batches, test subjects, ...
- one way of grouping runs (one blocking factor)



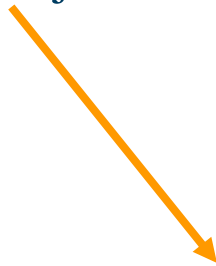
Split-plot experiments

- hard-to-change factors
 - levels are reset only a couple of times
 - gives rise to blocks of observations for which the hard-to-change factors are not reset
 - these blocks are called whole plots
- two-stage production processes
 - experiments often span more than one step in a production process
 - some factors are applied in a first stage
 - others are applied in a subsequent stage



Model

$$y_{ij} = \mathbf{f}'(\mathbf{x}_{ij})\boldsymbol{\beta} + \gamma_i + \varepsilon_{ij}$$



**experimental
factors**



Model

$$y_{ij} = \mathbf{f}'(\mathbf{x}_{ij})\boldsymbol{\beta} + \gamma_i + \varepsilon_{ij}$$

block effect

i.i.d. $N(0, \sigma_\gamma^2)$

**random
error**

i.i.d. $N(0, \sigma_\varepsilon^2)$



Model

$$y_{ij} = \mathbf{f}'(\mathbf{w}_i, \mathbf{s}_{ij})\boldsymbol{\beta} + \gamma_i + \varepsilon_{ij}$$

whole plot factors

sub-plot factors



Model

$$y_{ij} = \mathbf{f}'(\mathbf{w}_i, \mathbf{s}_{ij})\boldsymbol{\beta} + \gamma_i + \varepsilon_{ij}$$

**whole-plot
effect**

i.i.d. $N(0, \sigma_\gamma^2)$

**random
error**

i.i.d. $N(0, \sigma_\varepsilon^2)$



Model

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\gamma} + \boldsymbol{\varepsilon}$$

$$\mathbf{V} = \text{var}(\mathbf{Y})$$

$$= \sigma_{\varepsilon}^2 \mathbf{I}_n + \sigma_{\gamma}^2 \mathbf{Z}\mathbf{Z}'$$



Model

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$$\mathbf{V} = \text{var}(\mathbf{Y})$$

$$= \sigma_{\varepsilon}^2 \mathbf{I}_n + \sigma_{\gamma}^2 \mathbf{Z}\mathbf{Z}'$$

$$= \sigma_{\varepsilon}^2 (\mathbf{I}_n + \eta \mathbf{Z}\mathbf{Z}') \text{ where } \eta = \frac{\sigma_{\gamma}^2}{\sigma_{\varepsilon}^2}$$



Model

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\gamma} + \boldsymbol{\varepsilon}$$

$$\mathbf{V} = \text{var}(\mathbf{Y})$$

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$$= \sigma_{\varepsilon}^2 (\mathbf{I}_n + \eta \mathbf{Z}\mathbf{Z}') \text{ where } \eta = \frac{\sigma_{\gamma}^2}{\sigma_{\varepsilon}^2}$$

$$= \sigma_{\varepsilon}^2 \left(\mathbf{I}_n + \frac{\rho}{1-\rho} \mathbf{Z}\mathbf{Z}' \right) \text{ where } \rho = \frac{\sigma_{\gamma}^2}{\sigma_{\gamma}^2 + \sigma_{\varepsilon}^2}$$



Estimation

- generalized least squares regression

$$\hat{\beta} = (\mathbf{X}' \hat{\mathbf{V}}^{-1} \mathbf{X})^{-1} \mathbf{X}' \hat{\mathbf{V}}^{-1} \mathbf{Y}$$

- covariance matrix

$$\text{var}(\hat{\beta}) = (\mathbf{X}' \hat{\mathbf{V}}^{-1} \mathbf{X})^{-1}$$

- $\hat{\mathbf{V}}$ is the estimated covariance structure of the observations
- $\hat{\mathbf{V}}$ can be estimated using restricted maximum likelihood estimation (REML)



Design approaches

- split-plot designs:
 - stratum-by-stratum construction: Trinca & Gilmour (2001)
 - equivalent-estimation designs: Vining, Kowalski & Montgomery (2005), Parker, Kowalski & Vining (2006, 2007a, 2007b)
 - optimal design approach
- blocked experiments:
 - orthogonal blocking
 - optimal design approach



Optimal design

- blocked experiments: Goos & Vandebroek (2001)
- split-plot experiments: Goos & Vandebroek (2003), Jones & Goos (2007)
- D-optimality criterion: seeks designs that maximize determinant of information matrix on β

$$\mathbf{M} = \mathbf{X}' \mathbf{V}^{-1} \mathbf{X}$$



Weakness 1

8 runs

4 blocks of size 2

2 factors

full quadratic model

8 parameters

variance components
are not estimable !

Blocks	X_1	X_2
1	-1	1
	1	0
2	0	1
	-1	-1
3	0	-1
	1	1
4	1	-1
	-1	0



Weakness 2

- design depends on variance components through \mathbf{V}
- not on their absolute magnitude, but on their relative magnitude
- unelegant, but ...
 - in many cases, optimal design does not depend on η or ρ
 - in other cases, a design that is optimal for one value of η or ρ is highly efficient for other values
- setting $\eta = 1$ or $\rho = 0.5$ is recommended



When constructing optimal designs

1. take into account information matrix for variance components
2. use a prior distribution for either η or ρ



Variance components

- restricted maximum likelihood estimation



Variance components

- restricted maximum likelihood estimation
- expression for information matrix can be found in, for instance, Searle, Casella & McCulloch (1992):

$$\frac{1}{2} \begin{bmatrix} \text{tr}((\mathbf{PZZ}')^2) & \text{tr}(\mathbf{P}^2\mathbf{ZZ}') \\ \text{tr}(\mathbf{P}^2\mathbf{ZZ}') & \text{tr}(\mathbf{P}^2) \end{bmatrix}$$



Variance components

- restricted maximum likelihood estimation
- expression for information matrix can be found in, for instance, Searle, Casella & McCulloch (1992):

$$\frac{1}{2} \begin{bmatrix} \text{tr}((\mathbf{PZZ}')^2) & \text{tr}(\mathbf{P}^2\mathbf{ZZ}') \\ \text{tr}(\mathbf{P}^2\mathbf{ZZ}') & \text{tr}(\mathbf{P}^2) \end{bmatrix}$$

with

$$\mathbf{P} = \mathbf{V}^{-1} - \mathbf{V}^{-1}\mathbf{X}(\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}$$



Composite criterion

- find design that maximizes

$$D = |\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}|^{1/p} \times \left| \begin{array}{cc} \text{tr}((\mathbf{PZZ}')^2) & \text{tr}(\mathbf{P}^2\mathbf{ZZ}') \\ \text{tr}(\mathbf{P}^2\mathbf{ZZ}') & \text{tr}(\mathbf{P}^2) \end{array} \right|^{1/2}$$

with

$$\mathbf{P} = \mathbf{V}^{-1} - \mathbf{V}^{-1}\mathbf{X}(\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}$$

- depends on η or ρ too, so D should read $D(\eta)$ or $D(\rho)$



Bayesian composite criterion

- assume prior distribution for either η or ρ
- find design that maximizes

$$D_B = \int_0^{\infty} \log(D(\eta)) \pi(\eta) d\eta$$

or

$$D_B = \int_0^1 \log(D(\rho)) \pi(\rho) d\rho$$

with $\pi(\eta)$ and $\pi(\rho)$ the prior distributions



Bayesian composite criterion

- no analytical expressions for these integrals
- need to be evaluated numerically
- usually done using pseudo-Monte Carlo sampling
- for instance

$$D_B = \sum_{i=1}^R \log(D(\eta_i))$$

where η_i is a random draw from the prior distribution $\pi(\eta)$ and R is large



Lognormal prior for η

- Bayesian composite criterion

$$D_B = \int_0^{\infty} \log(D(\eta)) \frac{1}{\eta \sigma \sqrt{2\pi}} e^{-\{\log(\eta) - \mu\}^2 / 2\sigma^2} d\eta$$

- Gauss-Hermite quadrature
- Gaussian quadrature based on Stieltjes-Wigert polynomials

$$D_B \approx \sum_{i=1}^R w_i^{\text{GSW}} \log(D(e^{\mu} \alpha_i^{\text{GSW}}))$$

- R is small



Choice of μ and σ^2

- geometric mean: $m = e^\mu$
- geometric standard deviation: $s = e^\sigma$
- 99.7% of probability mass lies in $[m/s^3, ms^3]$
- if you expect η to be around one, choose a value of one for m (and, therefore, $\mu = 0$)
- if you expect η is in the interval $[0.1, 10]$, choose

$$s = \sqrt[3]{10} \text{ and } \sigma = \log(\sqrt[3]{10})$$





Beta prior for ρ

- Bayesian composite criterion

$$D_B = \int_0^1 \log(D(\rho)) \frac{\Gamma(\kappa + \lambda)}{\Gamma(\kappa)\Gamma(\lambda)} \rho^{\kappa-1} (1-\rho)^{\lambda-1} d\rho$$

- Gauss-Jacobi quadrature

$$D_B \approx \frac{1}{2^{\kappa+\lambda-1}} \cdot \frac{\Gamma(\kappa + \lambda)}{\Gamma(\kappa)\Gamma(\lambda)} \sum_{i=1}^R w_i^{\text{GJ}} \log \left(D \left(\frac{\alpha_i^{\text{GJ}} + 1}{2} \right) \right)$$


- R is small



Choice of κ and λ

- expectation of a beta distributed random variable is $\kappa / (\kappa + \lambda)$
- large values for the sum $\kappa + \lambda$ correspond to a smaller variance
- for the expectation $\kappa / (\kappa + \lambda)$ to be 0.5, κ and λ have to be identical
- in case of little prior uncertainty about the value of ρ , the sum $\kappa + \lambda$ has to be large



Example 1

8 runs

4 blocks of size 2

2 factors

full quadratic model

8 parameters



Example 1

Blocks	x_1	x_2
1	-1	1
	1	0
2	0	1
	-1	-1
3	0	-1
	1	1
4	1	-1
	-1	0

Old criterion

Blocks	x_1	x_2
1	1	1
	0	0
2	-1	-1
	-1	-1
3	1	-1
	-1	1
4	1	0
	0	1

New criterion



Example 1

(criterion values)

approach	composite criterion	fixed effects	variance components
any prior mean	-0.39	4.25	-2.2
old	$-\infty$	5.31	$-\infty$



Example 2

8 runs

4 whole plots of size 2

1 whole-plot factor (w)

1 sub-plot factor (s)

full quadratic model

8 parameters



Example 2

WPs	w	s
1	1	-1
	1	1
2	0	-1
	0	0
3	-1	-1
	-1	1
4	-1	1
	-1	0

Old criterion

WPs	w	s
1	1	1
	1	-1
2	-1	-1
	-1	1
3	-1	-1
	-1	1
4	0	-1
	0	0

**New criterion
(small prior mean)**

WPs	w	s
1	1	1
	1	-1
2	0	-1
	0	0
3	0	1
	0	0
4	-1	1
	-1	-1

**New criterion
(large prior mean)**



Example 2

(criterion values)

approach	composite criterion	fixed effects	variance components
small prior mean	-0.49	3.64	-2.20
large prior mean	-0.51	3.92	-2.33
old	-0.56	4.04	-2.46



Example 3

24 runs

6 blocks of size 4

5 factors

full quadratic model

23 parameters



Example 3

(average estimates)

η	approach	σ_ε^2	σ_Y^2	est η
10	new	0.985	11.732	10.612
	old	19.496	4.752	0.222
1	new	4.481	6.418	1.389
	old	14.355	2.824	0.220
0.1	new	6.820	1.755	0.301
	old	9.272	0.731	0.217



Example 3

(estimation problems)

η	appr.	zero σ_ε^2	large σ_ε^2	zero σ_γ^2	large σ_γ^2	conv.
10	new	531	21	319	8	1085
	old	753	77	3607	33	1571
1	new	240	12	1764	7	455
	old	874	85	4108	57	848
0.1	new	131	4	3701	1	116
	old	1053	84	4763	53	317



Discussion

- new design criterion with dual objective:
 - fixed-effects estimation
 - variance-component estimation
- Bayesian flavor
 - prior distribution for relative magnitude of variance components
 - quadrature approaches
- fewer estimation problems with variance components
- new approach is not a panacea: designs with few runs and blocks/whole plots will always be problematic



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