

Presentation at the Isaac Newton Institute, Cambridge

5 July 2005

*Comparisons of P-densities obtained from historical
asset prices, option prices and risk transformations*

Presented by Stephen Taylor

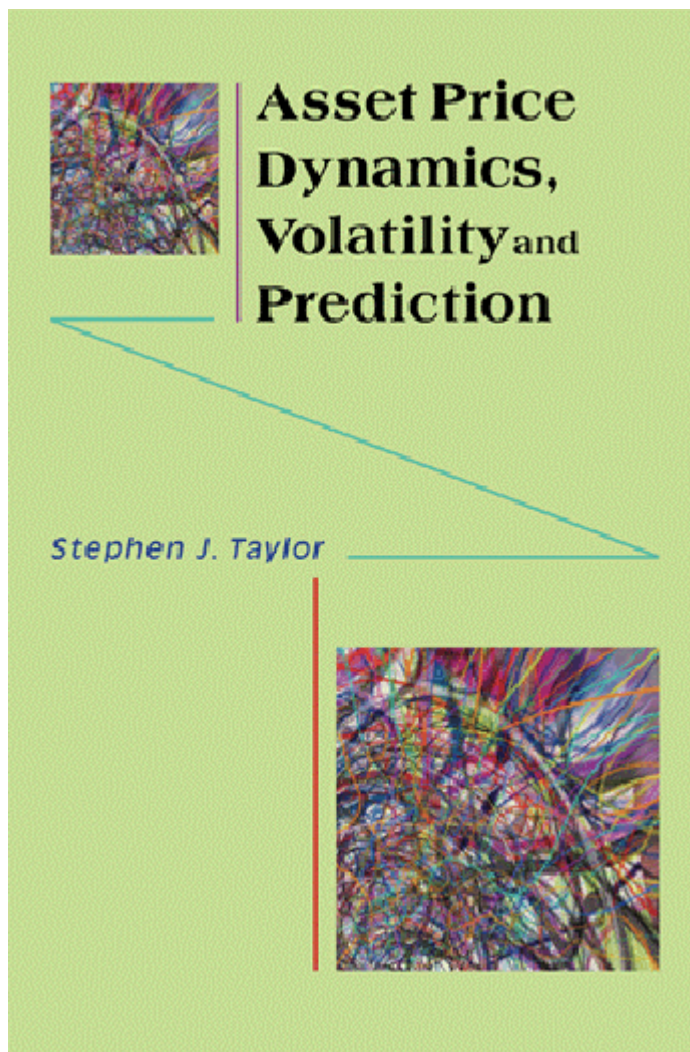
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Research commenced at:

Department of Accounting & Finance

Lancaster University, England



Available August/September 2005

For more information, including ordering, visit
www.pup.princeton.edu

Or

Use Google to search for *Asset Price Dynamics*

P & Q

Q refers to probabilities, densities, etc., for a *risk-neutral* economy.

P refers to the *real-world*.

Some people use other adjectives: true, objective, subjective, physical, historical.

As we see later, we can interconnect:

a P -density,

a Q -density,

a representative utility function.

Context

- Option prices certainly contain useful information about future volatility.
- Sets of option prices, for the same expiry time T and for different strikes X , provide implied densities.
- Can these *risk-neutral* densities be transformed to provide *real-world* densities, that are more informative than a time series of asset prices?
- Related research by Bliss & Panigirtzoglou and by Anagnou, Bedendo, Hodges & Tompkins.

Today's talk covers:

- Two papers
 - One is “under revision”, the other is preliminary
- Notation
- Theory
 - Risk-neutral densities
 - Transformations to real-world densities
 - Likelihood methods
- Data, initially U.K. and later U.S.
- Empirical results
 - Moments of densities
 - Calibration assessments
 - Likelihoods and incremental information
 - Measures of risk aversion

Study 1

Closed-form transformations from risk-neutral to real-world distributions

By: Liu, Shackleton, Taylor & Xu

PDF: at www.ssrn.com,
via <http://ssrn.com/author=44638>

Contributions

1. Closed-form real-world densities.
2. Likelihood methodology.

Applied to:

- Density estimation for the FTSE-100 index.
- With a four-week horizon.

Selected notation

S	The spot price now.
F	The futures price now.
T	The time until futures and options contracts cease trading.
$S_T = F_T$	Prices when trading of derivatives ends.
r	Risk-free interest rate.
X	Exercise (strike) price of an option.
$c(X)$	Price of a European call option.
σ	Volatility.
g	Risk-neutral density for S_T .
\tilde{g}	Real-world density for S_T .
G, \tilde{G}	Cumulative distribution functions.

Risk-neutral densities

A finite set of observed market prices $c_{market}(X_j)$ can not define a unique RND.

We estimate flexible densities using the spline method of Bliss & Panigirtzoglou.

More importantly, we estimate parametric RNDs from

$$\min_{\theta} \frac{1}{N} \sum_{j=1}^N (c_{market}(X_j) - c(X_j | \theta))^2,$$

with

$$c(X_j | \theta) = e^{-rT} \int_{X_j}^{\infty} (x - X_j) g(x | \theta) dx.$$

We consider two parametric families of RNDs.

A mixture of two lognormals

$$g_{MLN}(x|\theta) = wg_{LN}(x|F_1, \sigma_1, T) + (1-w)g_{LN}(x|F_2, \sigma_2, T)$$

with

$$g_{LN}(x|F, \sigma, T) = \frac{1}{x\sigma\sqrt{2\pi T}} \exp\left[-\frac{1}{2}\left(\frac{\log(x) - [\log(F) - \frac{1}{2}\sigma^2 T]}{\sigma\sqrt{T}}\right)^2\right]$$

Parameter vector $\theta = (F_1, F_2, \sigma_1, \sigma_2, w)$.

Risk-neutral constraint : $wF_1 + (1-w)F_2 = F$.

Generalised beta (GB2)

$$g_{GB2}(x|a, b, p, q) = \frac{a}{b^{ap} B(p, q)} \frac{x^{ap-1}}{\left[1 + \left(\frac{x}{b}\right)^a\right]^{p+q}}$$

Parameter vector $\theta = (a, b, p, q)$.

Risk-neutral constraint : $F = bB(p + a^{-1}, q - a^{-1})/B(p, q)$.

Two transformations to real-world densities (RWDs)

1. CRRA utility functions

The stochastic discount factor (pricing kernel) is

$$M(x) = e^{-rT} \frac{g(x)}{\tilde{g}(x)} .$$

When M is proportional to marginal utility, the RWD

for a power utility function, $v(x) \propto x^{1-\gamma}$, is

$$\tilde{g}(x) = \frac{x^\gamma g(x)}{\int_0^\infty y^\gamma g(y) dy} .$$

_____ g _____

_____ \tilde{g} _____

Lognormal

Lognormal, same σ

LN mixture

LN mixture, same σ_1, σ_2

GB2

GB2, same a, b

2. Statistical calibration

This method uses a calibration function C to transform the risk-neutral cumulative distribution function (c.d.f.), G into the real-world c.d.f. \tilde{G} , by

$$\tilde{G}(x) = C(G(x)).$$

The function C is a c.d.f.

Ideally:

- C is the real-world c.d.f. of $G(S_T)$,
- And then $\tilde{G}(S_T)$ has a uniform real-world distribution.

Following Fackler and King, we let the calibration function C be the Beta c.d.f. and then:

$$\tilde{g}(x) = \frac{G(x)^{j-1} (1-G(x))^{k-1}}{B(j, k)} g(x).$$

The transformation has parameters j and k .

- The RND & RWD coincide when $j = k = 1$.
- The implicit utility function is risk-averse if, and only if, $k \leq 1 \leq j$ with $j \neq k$.

Likelihood methods

Option expiry months, $i = 1, 2, 3, \dots$

Densities formed at times t_i .

Options expire at times t_i^* , at index levels $S_{T,i}$.

Non-overlapping densities, i.e. $t_i^* \leq t_{i+1}$.

Hence the likelihood of observed index levels is the product of densities, and

$$\log(L(S_{T,1}, S_{T,2}, \dots, S_{T,n} | \theta^*)) = \sum_{i=1}^n \log(\tilde{g}_i(S_{T,i} | \theta^*))$$

with θ^* any transformation parameter(s), γ or (j, k) , that can be obtained by maximising L .

Data

- Seek densities for the FTSE-100 index.
- Densities are defined 4 weeks before options expire.
- Historical real-world densities are obtained by estimating and simulating asymmetric ARCH models.
- Expiry months:
 - Original dataset has 83, June 1993 to April 2000.
 - New dataset has 126, ending December 2003.
- Risk-neutral densities are estimated:
 - Originally from *intraday midquote* and *trade* option prices,
 - For the new data, from option *settlement* prices.

Results

Risk-neutral densities

- Negatively skewed. Crash phobia!?
- Similar for GB2, lognormal mixtures and splines.
- Similar for midquote, trade and settlement prices.
- Exercise prices typically enclose 98% of the RND.

Transformation parameter estimates

- The CRRA parameter γ is near 4 for the shorter period but near 2 for the longer period – reflecting stock market falls after 2000.
- The calibration parameter j is near 1.40 (shorter period) and 1.25 (longer period), with k near 1.10.

Results are now for the original, shorter dataset

Moment comparisons

Of most interest for skewness. Median values are :

	<u>RND</u>	<u>RWD-u</u>	<u>RWD-c</u>
GB2	-0.75	-0.60	-0.56
Mix LN	-0.70	-0.59	-0.55

- Transformations reduce the level of skewness.
- Historical densities have median skewness -0.15 .
- So, option densities are much more skewed than the historical densities.
- Crash anxieties are unduly pessimistic!?

Calibration

Let $u_i = G(S_{T,i})$ or $\tilde{G}(S_{T,i})$.

Sample c.d.f. $\hat{C}(u) = \text{fraction of the } u_i \leq u$.

Good calibration requires $\hat{C}(u) \cong u$, all u .

But all four sets of RNDs have $\hat{C}(u) = 0$ for $u \leq 0.08$.

Historical and transformed densities are more satisfactorily calibrated.

The historical densities have the least value of

$$\max_u \left| \hat{C}(u) - u \right|.$$

Likelihoods

Maximum log-likelihood values for option methods,
relative to the historical densities, using midquote
 option prices:

	<u>RNDs</u>	<u>RWDs</u>	<u>RWDs</u>
<u>Transform...</u>		utility	calibration
<u>Method</u>			
Spline	-0.65	1.09	2.73
GB2	0.72	2.37	3.54
LN mixture	2.79	4.61	5.87

Likelihood-ratio tests

Option RNDs (null) versus option RWDs (alternative) :

- All eight test values reject the null at the 10% level
(midquote or trade prices, GB2 or LN mixture densities, either transformation).

Encompassing densities

$$\tilde{g}_{combined}(x|\alpha) = \alpha g_{RND/RWD}(x) + (1 - \alpha) \tilde{g}_{ARCH}(x)$$

- For RNDs, the average α is 0.57. All LR test values *reject* the null hypothesis $\alpha = 0$ at the 2% level.
- For RWDs, the average α is 0.71. All LR test values *accept* the null $\alpha = 1$ at the 5% level.

Empirical risk aversion

Within the representative agent framework, the utility function v satisfies

$$M(x) = \lambda \frac{dv}{dx} = e^{-rT} \frac{g(x)}{\tilde{g}(x)}.$$

Empirical estimates of the pricing kernel M :

- Use options data for g and historical for \tilde{g} ,
- Are shown as geometric means across all expiry months,
- Are generally decreasing functions of x , as required by theory ($v'' < 0$).
- Are hence unlike U.S. estimates.

Conclusions

For our U.K. data:

Option prices provide densities that are more informative than historical densities, *before* any risk adjustments are made.

Two simple transformations provide *real-world* densities from option prices that are more accurate than the *risk-neutral* densities.

Figure 1

The historical density and three risk-neutral densities, obtained from midquote option prices, on March 21, 1997

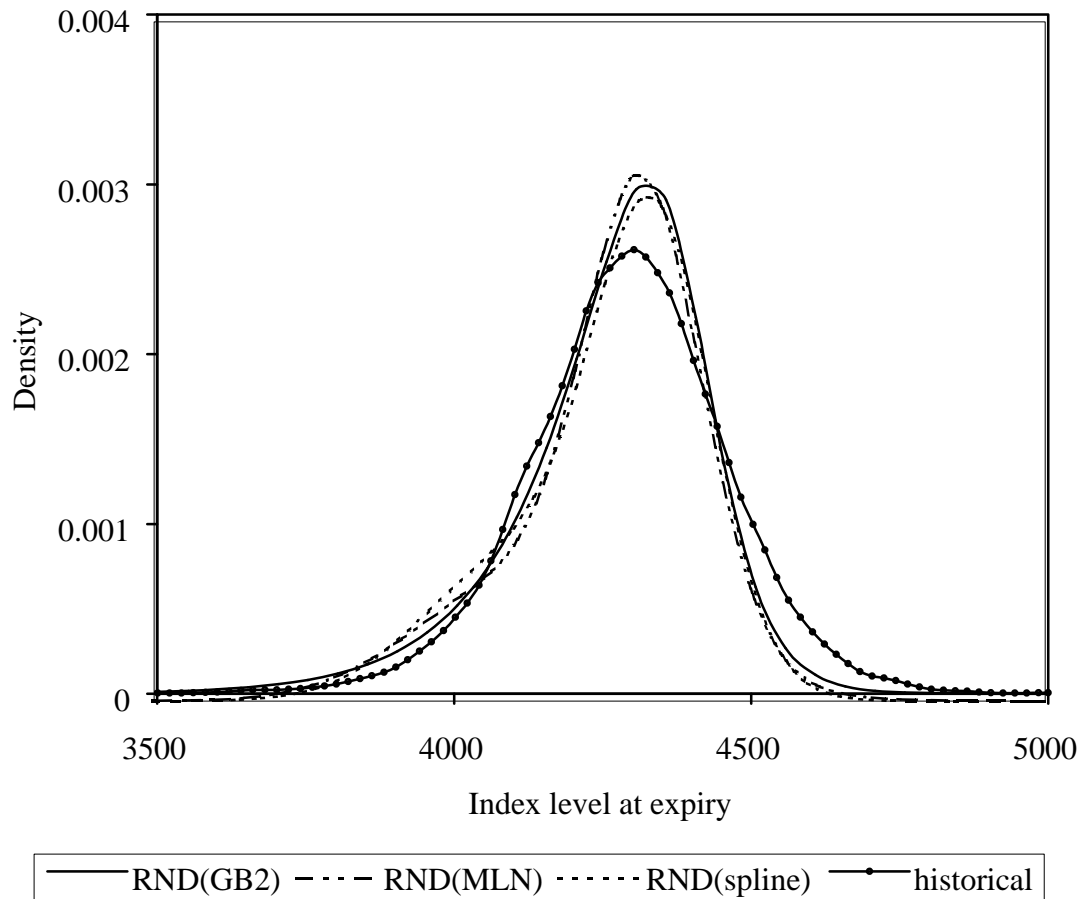


Figure 2

The risk-neutral GB2 density, from midquotes, and two corresponding real-world densities, on March 21, 1997

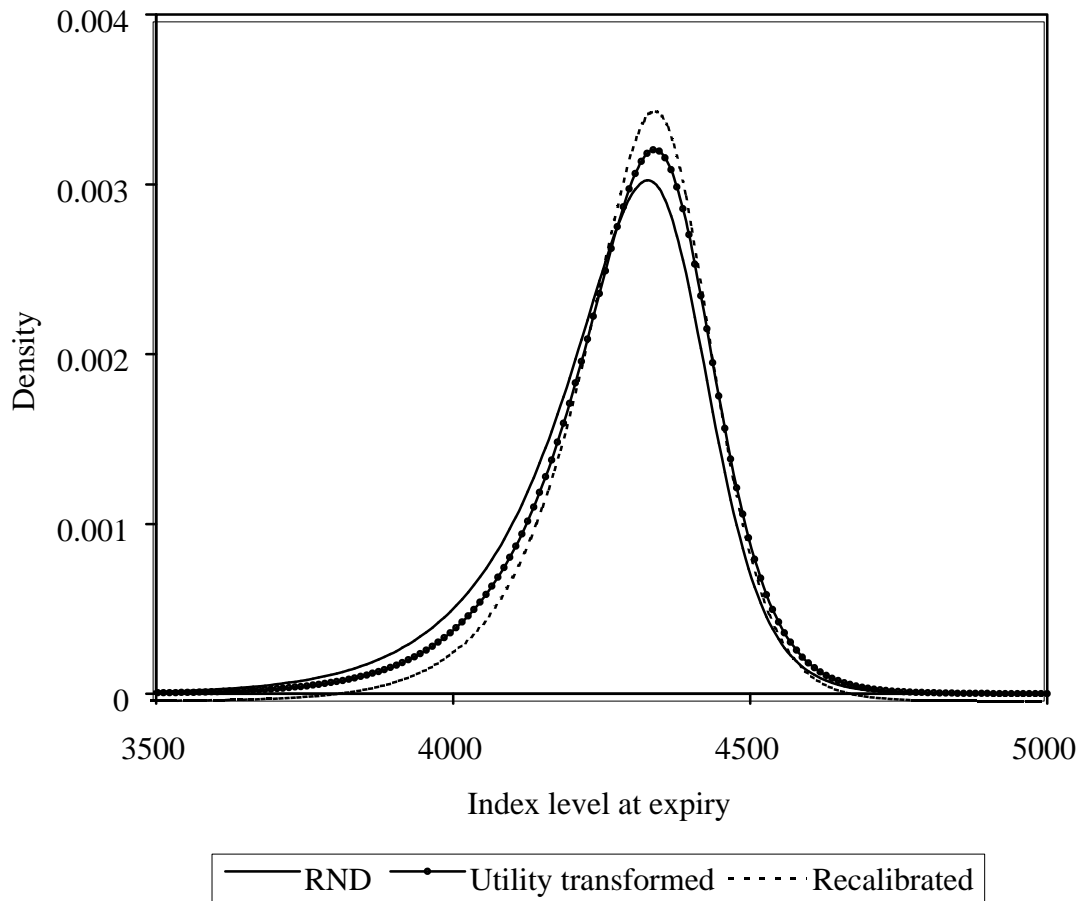


Figure 3

The risk-neutral lognormal mixture density, from midquotes, and two corresponding real-world densities, on March 21, 1997

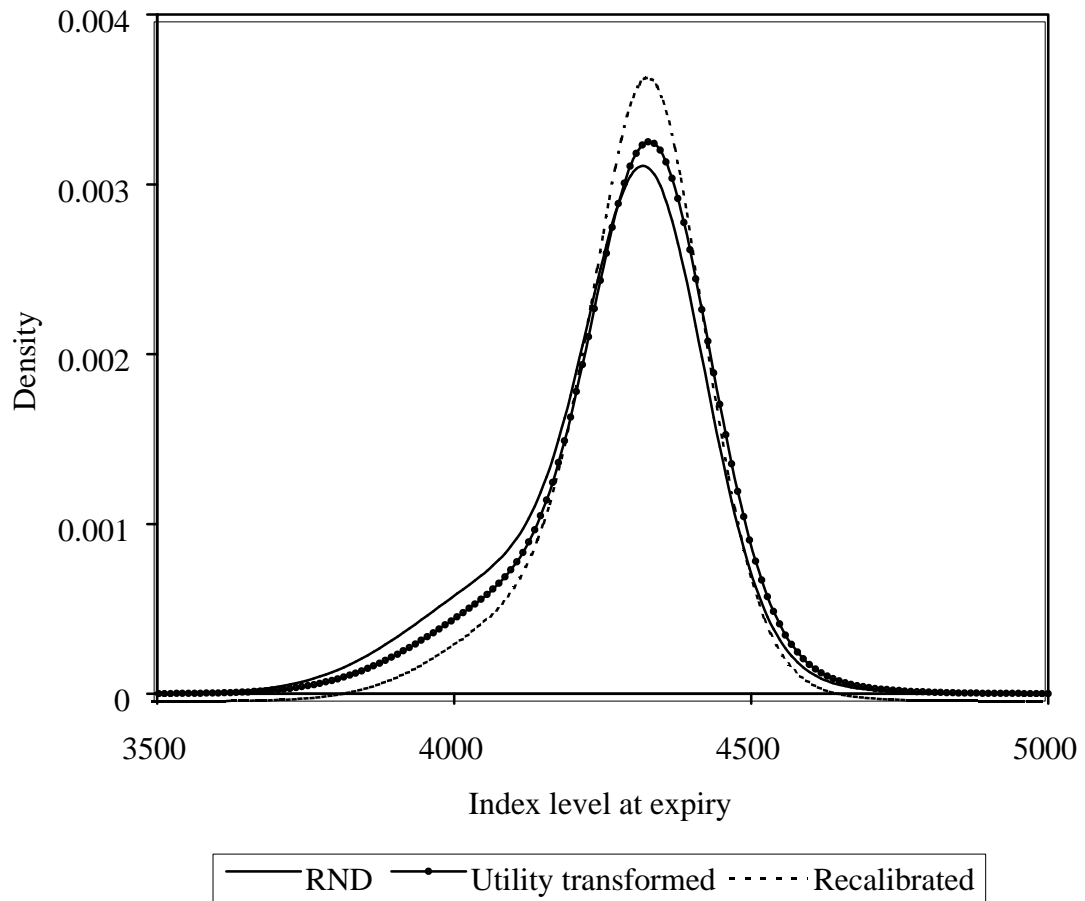


Figure 4

Cumulative distributions for cumulative probabilities obtained from GB2 RNDs and two sets of real-world densities

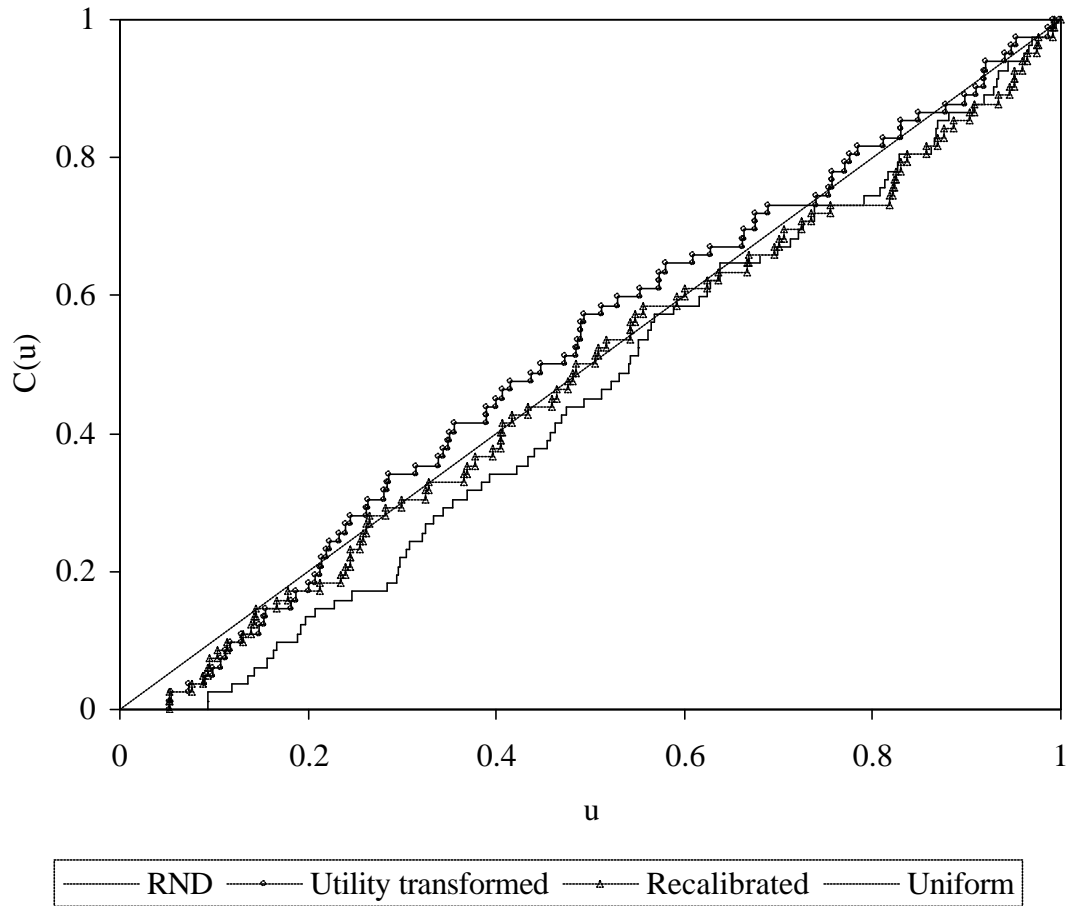


Figure 5

Cumulative distributions for cumulative probabilities obtained from historical densities

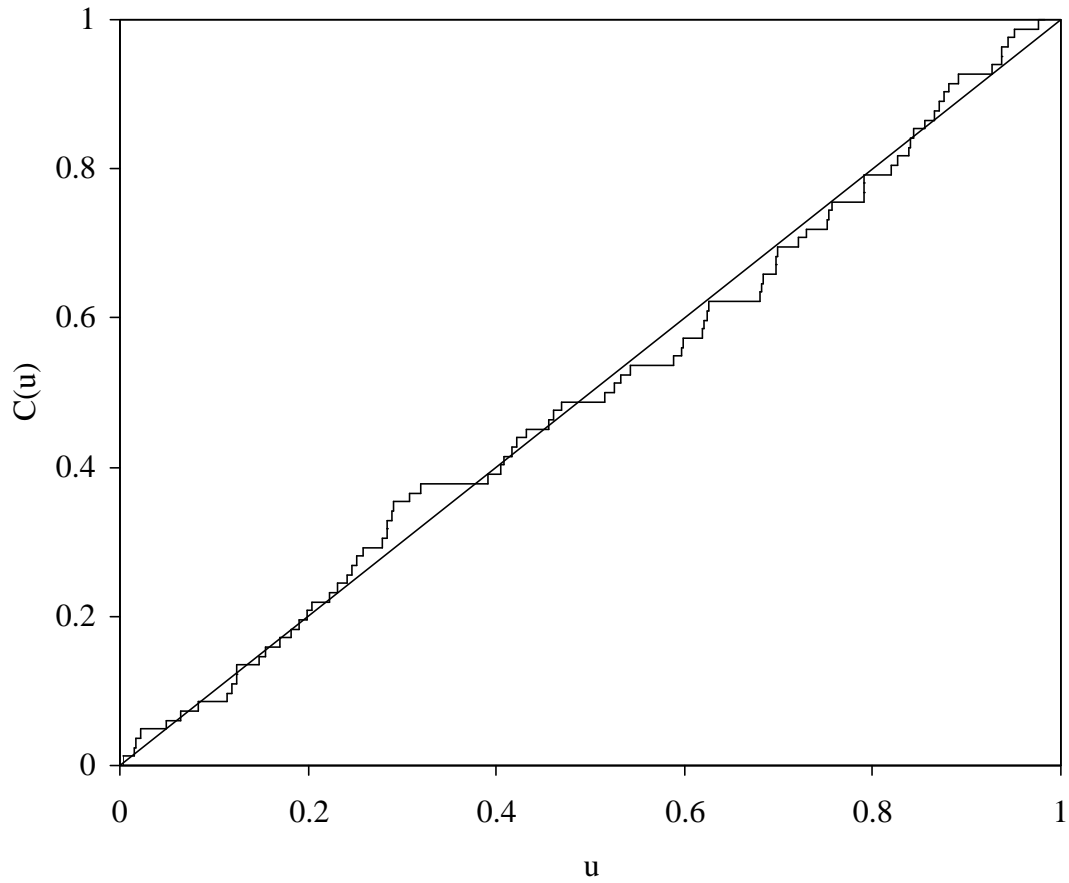
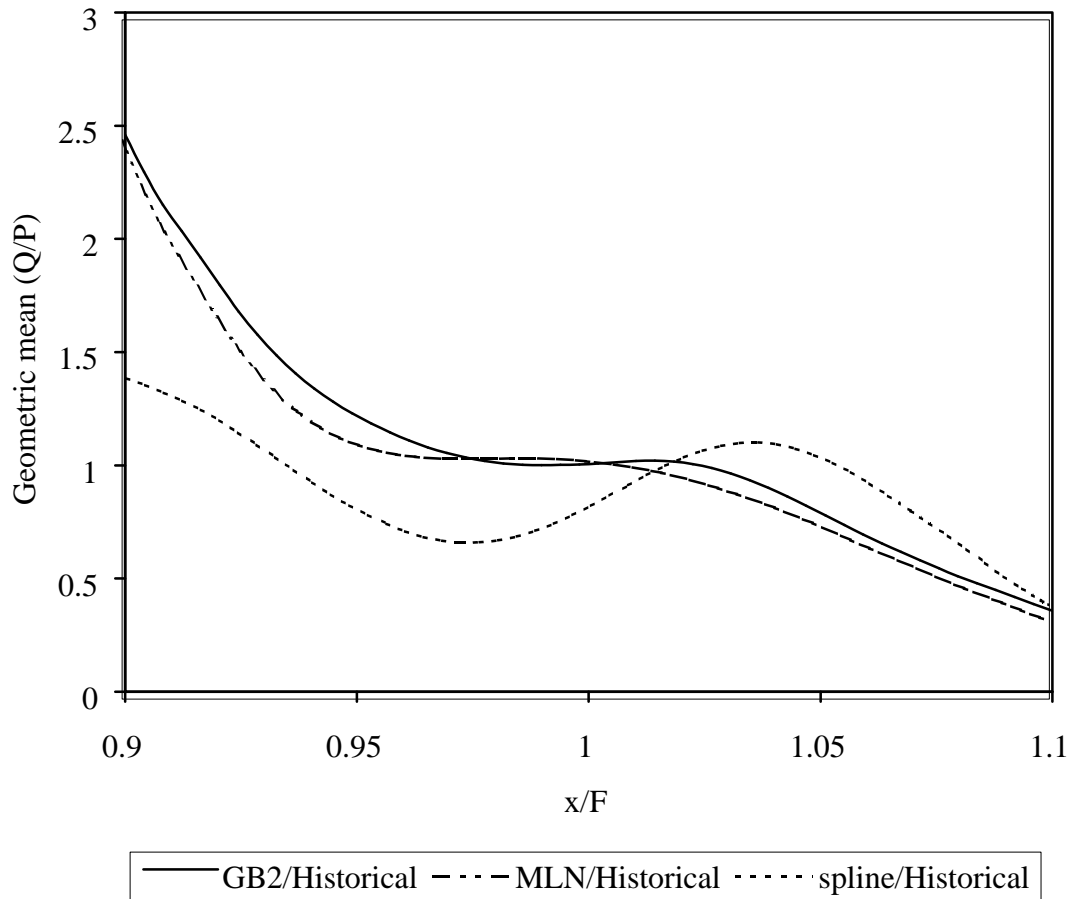


Figure 6

Empirical pricing kernels, averaged across all expiry months



Study 2

Density prediction for the S & P 500 index

By: Shackleton, Taylor & Yu

Preliminary: Paper to be completed in 2005

Contributions

1. Comparisons between densities estimated from option prices and from *high-frequency* index levels.
2. Densities obtained for *all horizons* from option prices.

Risk-neutral densities

Following Heston, suppose:

- the variance V follows a square-root process,
- the process for the futures prices, F , is continuous,
- the risk-neutral dynamics are:

$$dF = F\sqrt{V}dW,$$

$$dV = \kappa(\theta - V)dt + \xi\sqrt{V}dZ,$$

with

$$dWdZ = \rho dt.$$

Option prices can then be calculated from a “closed form” formula, that involves characteristic functions.

The initial variance V_0 and the parameters $\kappa, \theta, \xi, \rho$ can be estimated by matching a *matrix* of market prices $c_{market}(T_i, X_j)$ with theoretical Heston prices.

The risk-neutral density can then be deduced for any horizon T .

Data

- From 1990 to 2004.
- Daily and five-minute S & P 500 futures returns are used to estimate historical real-world densities from ARCH models. The data-frequency of observed returns for an ARCH model is matched with the density horizon.
- Risk-neutral densities are estimated from settlement prices for options on S & P 500 futures.
- Real-world densities, derived from option prices, are obtained by applying the calibration transformation mentioned earlier.

Horizons

1 day

1, 2, 4, 6, 8 & 12 weeks

Results (preliminary)

Likelihood comparisons

ARCH models versus Heston-RND

Horizon < 3 weeks: ARCH densities from high-frequency returns are best.

Horizon > 3 weeks: Heston-RND is best.

ARCH models versus calibrated Heston density

Horizon < 5 weeks: Calibrated Heston density best.

> 5 weeks: ARCH & calibrated densities are similar.

Encompassing densities

More weight is given to the calibrated Heston density than to the best ARCH density.

Diagnostic tests (5% significance level)

Kolmogorov-Smirnov test – All methods fail for the one-day horizon. Few failures at other horizons.

Berkowitz LR3 test –

- ARCH models usually pass.
- Heston RNDs fail for short horizons.
- Calibrated Heston densities pass for all horizons.